#### **Correlation Coefficient**

• 
$$r = \frac{\sum (x_i - \overline{x})(y_i - \overline{y})}{\sqrt{\sum (x_i - \overline{x})^2 \sum (y_i - \overline{y})^2}} = \frac{SSXY}{\sqrt{SSX \cdot SSY}}$$

recall that 
$$\hat{\beta}_1 = \frac{\sum (x_i - \overline{x})(y_i - \overline{y})}{\sum (x_i - \overline{x})^2} = \frac{SSXY}{SSX}$$

$$\hat{\beta}_1 = r \, \frac{S_Y}{S_X}$$

### Properties

$$\circ$$
  $-1 \le r \le 1$ 

 $\circ$  r does not depend on units of measurement (dimensionless)

 $\circ$   $r \approx 0 \rightarrow \text{no } linear \text{ relationship between } X \text{ and } Y$ 

 $\circ$  larger |r| means a stronger linear relationship

$$r^{2} = \frac{\sum (y_{i} - \overline{y})^{2} - \sum (y_{i} - \hat{y}_{i})^{2}}{\sum (y_{i} - \overline{y})^{2}} = \frac{SSY - SSE}{SSY}$$

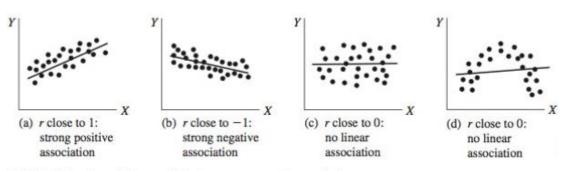
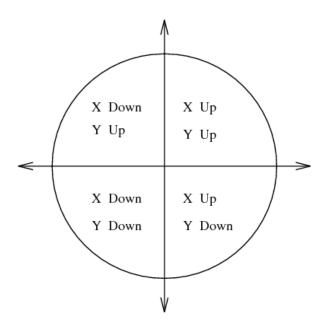


FIGURE 6.1 Correlation coefficient as a measure of association

# **Transformations and Tukey's "Rule of the Bulge"**

- Observe which way the curve bulges as suggested by a scatterplot of the data
- Transform y or x (or both) according to the signs of the corresponding quadrant:

up → powers >1down → powers <1 (including logarithm)</li>



### **Correlation Coefficient - Inference**

$$r = \frac{\sum (x_i - \overline{x})(y_i - \overline{y})}{\sqrt{\sum (x_i - \overline{x})^2 \sum (y_i - \overline{y})^2}} = \frac{SSXY}{\sqrt{SSX \cdot SSY}} \rightarrow \hat{\beta}_1 = r \frac{S_Y}{S_X}$$

- Testing Hypotheses about  $\rho$ , the *population* correlation coefficient
- $H_0: \rho = 0$ 
  - Sampling distribution of *r* is symmetric & approx. normal
  - $t = \frac{r\sqrt{n-2}}{\sqrt{1-r^2}} \text{ with } n-2 \text{ d.f. is equivalent to testing H}_0: \beta_1=0 \text{ with } t = \frac{\hat{\beta}_1 0}{SE(\hat{\beta}_1)}$
- $H_0$ :  $\rho = \rho_0$  where  $\rho_0 \neq 0$ 
  - $\circ$  Sampling distribution of r is NOT symmetric or approx. normal
  - o Fisher's "Z-transformation" gives an approx. normally distributed statistic

$$\frac{1}{2} \cdot \ln \left( \frac{1+r}{1-r} \right) \sim N \left( \frac{1}{2} \cdot \ln \left( \frac{1+\rho}{1-\rho} \right), \frac{1}{n-3} \right)$$

o Testing H<sub>0</sub>: 
$$\rho = \rho_0$$
  $Z = \frac{\frac{1}{2} \cdot \ln(1 + r/1 - r) - \frac{1}{2} \cdot \ln(1 + \rho_0/1 - \rho_0)}{1/\sqrt{n-3}} \sim N(0,1)$ 

$$0 \quad 100(1-\alpha)\% \text{ CI for } \frac{1}{2} \cdot \ln\left(\frac{1+\rho}{1-\rho}\right) : \qquad \frac{1}{2} \cdot \ln\left(\frac{1+r}{1-r}\right) \pm z_{\alpha/2} \cdot \frac{1}{\sqrt{n-3}} = (L_z, U_z)$$

- ο 100(1-α)% CI for  $\rho$ : ( $L_{\rho}$ ,  $U_{\rho}$ ) = (lower endpoint, upper endpoint)
  - Find  $L_z \& U_z$  above and solve for  $L_\rho \& U_\rho$  in  $L_z = \frac{1}{2} \cdot \ln \left( \frac{1 + L_\rho}{1 L_\rho} \right)$   $U_z = \frac{1}{2} \cdot \ln \left( \frac{1 + U_\rho}{1 U_\rho} \right)$

•  $H_0$ :  $\rho_1 = \rho_2$  (for 2 *independent* samples)

o let 
$$W_1 = \frac{1}{2} \cdot \ln \left( \frac{1 + r_1}{1 - r_1} \right)$$
 and  $W_2 = \frac{1}{2} \cdot \ln \left( \frac{1 + r_2}{1 - r_2} \right)$ 

$$Z = \frac{W_1 - W_2}{\sqrt{1/(n_1 - 3) + 1/(n_2 - 3)}} \sim N(0, 1)$$

## **Bivariate Normal Distribution**

$$f(x,y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}}e^{\frac{-1}{2(1-\rho^2)}\left(\left(\frac{x-\mu_x}{\sigma_x}\right)^2 - 2\rho\left(\frac{x-\mu_x}{\sigma_x}\right)\left(\frac{y-\mu_y}{\sigma_y}\right) + \left(\frac{y-\mu_y}{\sigma_y}\right)^2\right)}$$

$$E(Y \mid X = x) = \mu_y + \rho \frac{\sigma_y}{\sigma_x} (x - \mu_x)$$

$$\hat{\mu}_{Y \mid X = x} = \overline{y} + r \frac{s_y}{s_x} (x - \overline{x})$$

