Lognormals and friends

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Computational Story Lab | Vermont Complex Systems Center Vermont Advanced Computing Core | University of Vermont























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Lognormals

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Outline

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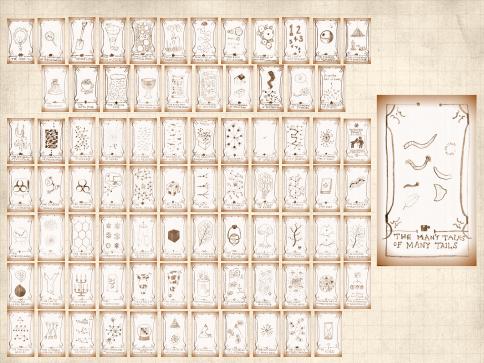
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Alternative distributions

There are other 'heavy-tailed' distributions:

1. The Log-normal distribution ☑

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \mathrm{exp}\left(-\frac{(\ln\!x - \mu)^2}{2\sigma^2}\right)$$

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$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln\!x - \mu)^2}{2\sigma^2}\right)$$

2. Weibull distributions

$$P(x)dx = \frac{k}{\lambda} \left(\frac{x}{\lambda}\right)^{\mu - 1} e^{-(x/\lambda)^{\mu}} dx$$

CCDF = stretched exponential ♂.

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3. Also: Gamma distribution , Erlang distribution , and more.

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The lognormal distribution:

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln\!x - \mu)^2}{2\sigma^2}\right)$$

- \Re lnx is distributed according to a normal distribution with mean μ and variance σ .
- Appears in economics and biology where growth increments are distributed normally.

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 \clubsuit Standard form reveals the mean μ and variance σ^2 of the underlying normal distribution:

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Lognormals **Empirical Confusability**

Growth Model



& Standard form reveals the mean μ and variance σ^2 of the underlying normal distribution:

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right)$$

For lognormals:

$$\begin{split} \mu_{\rm lognormal} &= e^{\mu + \frac{1}{2}\sigma^2}, \qquad {\rm median}_{\rm lognormal} = e^{\mu}, \\ \sigma_{\rm lognormal} &= (e^{\sigma^2} - 1)e^{2\mu + \sigma^2}, \qquad {\rm mode}_{\rm lognormal} = e^{\mu - \sigma^2}. \end{split}$$

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All moments of lognormals are finite.

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Take *Y* as distributed normally:

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Take *Y* as distributed normally:



$$P(y) \mathrm{d}y \, = \frac{1}{\sqrt{2\pi}\sigma} \mathrm{exp}\left(-\frac{(y-\mu)^2}{2\sigma^2}\right) \mathrm{d}y$$

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$$P(y) \mathrm{d} y \, = \frac{1}{\sqrt{2\pi}\sigma} \mathrm{exp} \left(-\frac{(y-\mu)^2}{2\sigma^2} \right) \mathrm{d} y$$

Set $Y = \ln X$:

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Set $Y = \ln X$:



 \Longrightarrow Transform according to P(x)dx = P(y)dy:



$$\frac{\mathrm{d}y}{\mathrm{d}x} = 1/x \Rightarrow \mathrm{d}y = \mathrm{d}x/x$$

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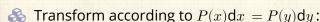


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Set $Y = \ln X$:





$$\frac{\mathrm{d}y}{\mathrm{d}x} = 1/x \Rightarrow \mathrm{d}y = \mathrm{d}x/x$$



$$\Rightarrow P(x) \mathrm{d}x = \frac{1}{\textcolor{red}{x}\sqrt{2\pi}\sigma} \mathrm{exp}\left(-\frac{(\ln \hspace{-.05cm}{x} - \mu)^2}{2\sigma^2}\right) \mathrm{d}x$$

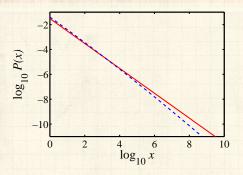
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Confusion between lognormals and pure power laws



Near agreement over four orders of magnitude!

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Empirical Confusability Growth Model

References



 \clubsuit For lognormal (blue), $\mu = 0$ and $\sigma = 10$.



 \clubsuit For power law (red), $\gamma = 1$ and c = 0.03.



What's happening:

$$\ln\!P(x) = \ln\left\{\frac{1}{x\sqrt{2\pi}\sigma}\!\exp\left(-\frac{(\ln\!x-\mu)^2}{2\sigma^2}\right)\right\}$$

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What's happening:

$$\begin{split} \ln\!P(x) &= \ln\left\{\frac{1}{x\sqrt{2\pi}\sigma}\!\exp\left(-\frac{(\ln\!x-\mu)^2}{2\sigma^2}\right)\right\} \\ &= -\!\ln\!x - \!\ln\!\sqrt{2\pi}\sigma - \frac{(\ln\!x-\mu)^2}{2\sigma^2} \end{split}$$

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$$= -\frac{1}{2\sigma^2} (\ln x)^2 + \left(\frac{\mu}{\sigma^2} - 1\right) \ln x - \ln \sqrt{2\pi}\sigma - \frac{\mu^2}{2\sigma^2}.$$

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$$=-\frac{1}{2\sigma^2}(\ln\!x)^2+\left(\frac{\mu}{\sigma^2}-1\right)\ln\!x-\ln\!\sqrt{2\pi}\sigma-\frac{\mu^2}{2\sigma^2}.$$

If the first term is relatively small,

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If the first term is relatively small,

$$\ln P(x) \sim -\left(1 - \frac{\mu}{\sigma^2}\right) \ln x + {
m const.}$$

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If the first term is relatively small,

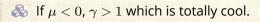
$$\boxed{ \ln\! P(x) \sim - \left(1 - \frac{\mu}{\sigma^2}\right) \ln\! x + \mathrm{const.} } \Rrightarrow \boxed{ \gamma = 1 - \frac{\mu}{\sigma^2} }$$

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Random Growth with Variable Lifespan



 \clubsuit If $\mu < 0$, $\gamma > 1$ which is totally cool.

 $\mbox{\&}\ \ \mbox{If } \mu>0 \mbox{, } \gamma<1 \mbox{, not so much.}$

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 \clubsuit If $\mu < 0$, $\gamma > 1$ which is totally cool.

 $\red{\$}$ If $\sigma^2\gg 1$ and μ ,

 $\ln P(x) \sim -\ln x + \mathrm{const.}$

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$$\ln P(x) \sim -\ln x + \text{const.}$$

Expect -1 scaling to hold until $(\ln x)^2$ term becomes significant compared to $(\ln x)$:

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$$-rac{1}{2\sigma^2}({
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$$\begin{split} &-\frac{1}{2\sigma^2}(\ln\!x)^2 \simeq 0.05 \left(\frac{\mu}{\sigma^2}-1\right) \ln\!x \\ \Rightarrow &\log_{10}\!x \lesssim 0.05 \times 2(\sigma^2-\mu) \!\log_{10}\!e \end{split}$$

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⇒ If you find a -1 exponent, you may have a lognormal distribution... PoCS, Vol. 1 Lognormals and friends 13 of 26

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Generating lognormals:

Random multiplicative growth:



$$x_{n+1} = rx_n$$

where r>0 is a random growth variable

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Random multiplicative growth:



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(Shrinkage is allowed)

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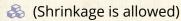


Random multiplicative growth:



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In log space, growth is by addition:

$$\ln x_{n+1} = \ln r + \ln x_n$$

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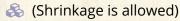


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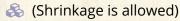


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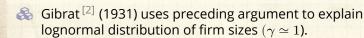
 $\Leftrightarrow \ln x_n$ is normally distributed

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- Gibrat [2] (1931) uses preceding argument to explain lognormal distribution of firm sizes ($\gamma \simeq 1$).
- & But Robert Axtell [1] (2001) shows a power law fits the data very well with $\gamma=2$, not $\gamma=1$ (!)

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- Problem of data censusing (missing small firms).

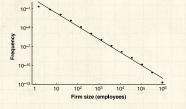
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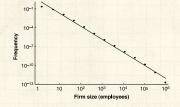
Freq $\propto (\text{size})^{-\gamma}$ $\gamma \simeq 2$ PoCS, Vol. 1 Lognormals and friends 16 of 26

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Freq $\propto (\text{size})^{-\gamma}$ $\gamma \simeq 2$

One piece in Gibrat's model seems okay empirically: Growth rate r appears to be independent of firm size. [1].

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Axtel cites Malcai et al.'s (1999) argument [5] for why power laws appear with exponent $\gamma \simeq 2$

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 \clubsuit The set up: N entities with size $x_i(t)$

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 $\red {\Bbb R}$ The set up: N entities with size $x_i(t)$

<page-header> Generally:

$$x_i(t+1) = rx_i(t)$$

where r is drawn from some happy distribution

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备 Generally:

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Same as for lognormal but one extra piece.

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- $\ref{eq:set}$ The set up: N entities with size $x_i(t)$
- 🚳 Generally:

$$x_i(t+1) = rx_i(t)$$

where r is drawn from some happy distribution

- Same as for lognormal but one extra piece.
- $\ensuremath{\&}$ Each x_i cannot drop too low with respect to the other sizes:

$$x_i(t+1) = \max(rx_i(t), c\left\langle x_i \right\rangle)$$

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Some math later...
Insert question from assignment 7 🗗

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Insert question from assignment 7 🗷



Find $P(x) \sim x^{-\gamma}$

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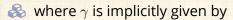
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Insert question from assignment 7 🗷



Find
$$P(x) \sim x^{-\gamma}$$



$$N = \frac{(\gamma - 2)}{(\gamma - 1)} \left[\frac{(c/N)^{\gamma - 1} - 1}{(c/N)^{\gamma - 1} - (c/N)} \right]$$

N = total number of firms.

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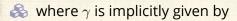
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Insert question from assignment 7 2



Find
$$P(x) \sim x^{-\gamma}$$



$$N = \frac{(\gamma - 2)}{(\gamma - 1)} \left[\frac{(c/N)^{\gamma - 1} - 1}{(c/N)^{\gamma - 1} - (c/N)} \right]$$

N = total number of firms.



Now, if
$$c/N \ll 1$$
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Insert question from assignment 7 2



Find
$$P(x) \sim x^{-\gamma}$$

$$N = \frac{(\gamma - 2)}{(\gamma - 1)} \left[\frac{(c/N)^{\gamma - 1} - 1}{(c/N)^{\gamma - 1} - (c/N)} \right]$$

N = total number of firms.



Now, if
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Which gives
$$\gamma \sim 1 + \frac{1}{1-c}$$

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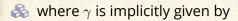
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Insert question from assignment 7 🗷



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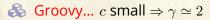
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Ages of firms/people/... may not be the same

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Ages of firms/people/... may not be the same



 \mathbb{A} Allow the number of updates for each size x_i to vary

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Ages of firms/people/... may not be the same

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 \Longrightarrow Example: $P(t)dt = ae^{-at}dt$ where t = age.

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Ages of firms/people/... may not be the same

 \Leftrightarrow Example: $P(t)dt = ae^{-at}dt$ where t = age.

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- Sizes are distributed as [6]

$$P(x) = \int_{t=0}^{\infty} a e^{-at} \frac{1}{x\sqrt{2\pi t}} \exp\left(-\frac{(\ln x - \mu)^2}{2t}\right) \mathrm{d}t$$

(Assume for this example that $\sigma \sim t$ and $\mu = \ln m$)

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Now averaging different lognormal distributions.

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Averaging lognormals



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Averaging lognormals



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🙈 Insert fabulous calculation (team is spared).

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Insert fabulous calculation (team is spared).

Some enjoyable suffering leads to:

$$P(x) \propto x^{-1} e^{-\sqrt{2\lambda(\ln\frac{x}{m})^2}}$$

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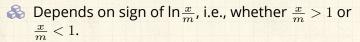
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$$P(x) \propto x^{-1} e^{-\sqrt{2\lambda(\ln\frac{x}{m})^2}}$$

 \clubsuit Depends on sign of $\ln \frac{x}{m}$, i.e., whether $\frac{x}{m} > 1$ or $\frac{x}{m} < 1.$



$$P(x) \propto \left\{ \begin{array}{ll} x^{-1+\sqrt{2\lambda}} & \text{if } \frac{x}{m} < 1 \\ x^{-1-\sqrt{2\lambda}} & \text{if } \frac{x}{m} > 1 \end{array} \right.$$

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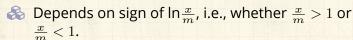
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'Break' in scaling (not uncommon)

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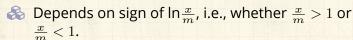
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- 'Break' in scaling (not uncommon)
- 🙈 Double-Pareto distribution 🗹

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- 'Break' in scaling (not uncommon)
- Double-Pareto distribution
- First noticed by Montroll and Shlesinger [7, 8]

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 ✓
- First noticed by Montroll and Shlesinger [7, 8]
- Later: Huberman and Adamic [3, 4]: Number of pages per website

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& Lognormals and power laws can be awfully similar



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🚵 Lognormals and power laws can be awfully similar

Random Multiplicative Growth leads to lognormal distributions



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Enforcing a minimum size leads to a power law tail



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🙈 Lognormals and power laws can be awfully similar

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With no minimum size but a distribution of lifetimes, the double Pareto distribution appears



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🙈 Enforcing a minimum size leads to a power law tail

With no minimum size but a distribution of lifetimes, the double Pareto distribution appears

Take-home message: Be careful out there...



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