Lognormals and friends

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Prof. Peter Sheridan Dodds | @peterdodds

Computational Story Lab | Vermont Complex Systems Center Vermont Advanced Computing Core | University of Vermont



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References



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200 1 of 26

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References





200 3 of 26

Outline

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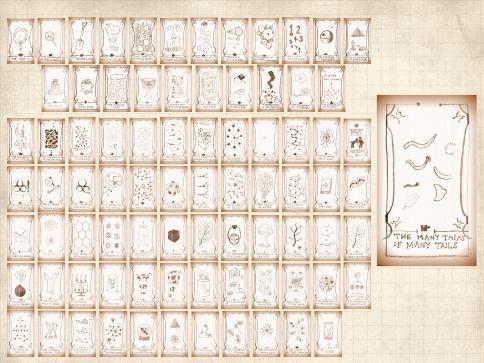
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References





200 4 of 26



Alternative distributions

There are other 'heavy-tailed' distributions:1. The Log-normal distribution 了

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right)$$

2. Weibull distributions

$$P(x)dx = \frac{k}{\lambda} \left(\frac{x}{\lambda}\right)^{\mu-1} e^{-(x/\lambda)^{\mu}} dx$$

CCDF = stretched exponential C.
Also: Gamma distribution C, Erlang distribution C, and more.



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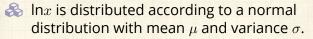
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The lognormal distribution:

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right)$$



Appears in economics and biology where growth increments are distributed normally.



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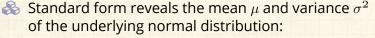
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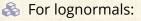
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$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right)$$



 $\mu_{ ext{lognormal}} = e^{\mu + rac{1}{2}\sigma^2}, \qquad ext{median}_{ ext{lognormal}} = e^{\mu},$

 $\sigma_{\rm lognormal} = (e^{\sigma^2}-1)e^{2\mu+\sigma^2}, \qquad {\rm mode}_{\rm lognormal} = e^{\mu-\sigma^2}.$

All moments of lognormals are finite.





Derivation from a normal distribution Take *Y* as distributed normally:

2

$$P(y)dy = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{(y-\mu)^2}{2\sigma^2}\right) dy$$

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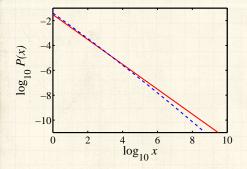
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Set $Y = \ln X$: Transform according to P(x)dx = P(y)dy: $\frac{dy}{dx} = 1/x \Rightarrow dy = dx/x$ $\Rightarrow P(x)dx = \frac{1}{x\sqrt{2\pi\sigma}} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) dx$



Confusion between lognormals and pure power laws



Near agreement over four orders of magnitude! PoCS, Vol. 1 @pocsvox

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So For lognormal (blue), $\mu = 0$ and $\sigma = 10$. For power law (red), $\gamma = 1$ and c = 0.03.



200 11 of 26

Confusion

What's happening:

$$\ln P(x) = \ln \left\{ \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) \right\}$$

$$= -\ln x - \ln \sqrt{2\pi}\sigma - \frac{(\ln x - \mu)^2}{2\sigma^2}$$

$$= -\frac{1}{2\sigma^2}(\ln x)^2 + \left(\frac{\mu}{\sigma^2} - 1\right)\ln x - \ln\sqrt{2\pi}\sigma - \frac{\mu^2}{2\sigma^2}.$$

If the first term is relatively small,

$$\boxed{\ln P(x) \sim -\left(1 - \frac{\mu}{\sigma^2}\right)\ln x + \text{const.}} \Rrightarrow \boxed{\gamma = 1 - \frac{\mu}{\sigma^2}}$$

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Confusion

If \$\mu < 0\$, \$\gamma > 1\$ which is totally cool.
If \$\mu > 0\$, \$\gamma < 1\$, not so much.
If \$\sigma^2 > 1\$ and \$\mu\$,

 $\ln P(x) \sim -\ln x + \text{const.}$

Solution Expect -1 scaling to hold until $(\ln x)^2$ term becomes significant compared to $(\ln x)$:

 $-\frac{1}{2\sigma^2}(\ln x)^2 \simeq 0.05 \left(\frac{\mu}{\sigma^2} - 1\right) \ln x$

$$\Rightarrow \log_{10} x \lesssim 0.05 \times 2(\sigma^2 - \mu) \log_{10} e \simeq 0.05(\sigma^2 - \mu)$$

♣ ⇒ If you find a -1 exponent, you may have a lognormal distribution... PoCS, Vol. 1 @pocsvox

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Generating lognormals:

3

Random multiplicative growth:

 $x_{n+1} = rx_n$

where r > 0 is a random growth variable (Shrinkage is allowed) (Shrinkage, growth is by addition:

 $\ln x_{n+1} = \ln r + \ln x_n$

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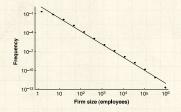
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Lognormals or power laws?

- Gibrat ^[2] (1931) uses preceding argument to explain lognormal distribution of firm sizes ($\gamma \simeq 1$).
- But Robert Axtell^[1] (2001) shows a power law fits the data very well with $\gamma = 2$, not $\gamma = 1$ (!)
- Problem of data censusing (missing small firms).



 $\begin{array}{l} {\rm Freq} \propto ({\rm size})^{-\gamma} \\ \gamma \simeq 2 \end{array}$

One piece in Gibrat's model seems okay empirically: Growth rate *r* appears to be independent of firm size.^[1]. PoCS, Vol. 1 @pocsvox

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An explanation

Axtel cites Malcai et al.'s (1999) argument ^[5] for why power laws appear with exponent $\gamma \simeq 2$ The set up: N entities with size $x_i(t)$ Generally:

 $x_i(t+1) = r x_i(t)$

where r is drawn from some happy distribution
Same as for lognormal but one extra piece.
Each x_i cannot drop too low with respect to the other sizes:

$$x_i(t+1) = \max(rx_i(t), c\left< x_i \right>)$$

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Some math later...

3

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2

Insert question from assignment 7 🖸

Find
$$P(x) \sim x^{-\gamma}$$

 \circledast where γ is implicitly given by

$$N = \frac{(\gamma - 2)}{(\gamma - 1)} \left[\frac{(c/N)^{\gamma - 1} - 1}{(c/N)^{\gamma - 1} - (c/N)} \right]$$

N = total number of firms.

Now, if
$$c/N \ll 1$$
 and $\gamma > 2$ $N = \frac{(\gamma - 2)}{(\gamma - 1)} \left[\frac{-1}{-(c/N)} \right]$

Which gives
$$\gamma \sim 1 + \frac{1}{1-c}$$

 \clubsuit Groovy... $c \text{ small} \Rightarrow \gamma \simeq 2$

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The second tweak

Ages of firms/people/... may not be the same

- \clubsuit Allow the number of updates for each size \boldsymbol{x}_i to vary
 - S Example: $P(t)dt = ae^{-at}dt$ where t = age.
- Reack to no bottom limit: each x_i follows a lognormal

🚳 Sizes are distributed as [6]

$$P(x) = \int_{t=0}^{\infty} a e^{-at} \frac{1}{x\sqrt{2\pi t}} \exp\left(-\frac{(\ln x - \mu)^2}{2t}\right) \mathrm{d}t$$

(Assume for this example that $\sigma \sim t$ and $\mu = \ln m$) Now averaging different lognormal distributions. PoCS, Vol. 1 @pocsvox

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Averaging lognormals

2

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Insert fabulous calculation (team is spared).
Some enjoyable suffering leads to:

 $P(x) \propto x^{-1} e^{-\sqrt{2\lambda(\ln \frac{x}{m})^2}}$

 $P(x) = \int_{t=0}^{\infty} ae^{-at} \frac{1}{x\sqrt{2\pi t}} \exp\left(-\frac{(\ln\frac{x}{m})^2}{2t}\right) dt$



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nac 21 of 26

The second tweak

2

a.

$$P(x) \propto x^{-1} e^{-\sqrt{2\lambda (\ln \frac{x}{m})^2}}$$

Solution Depends on sign of $\ln \frac{x}{m}$, i.e., whether $\frac{x}{m} > 1$ or $\frac{x}{m} < 1$.

$$P(x) \propto \begin{cases} x^{-1+\sqrt{2\lambda}} & \text{if } \frac{x}{m} < 1\\ x^{-1-\sqrt{2\lambda}} & \text{if } \frac{x}{m} > 1 \end{cases}$$

Break' in scaling (not uncommon)
 Double-Pareto distribution
 First noticed by Montroll and Shlesinger ^[7, 8]
 Later: Huberman and Adamic ^[3, 4]: Number of pages per website

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Summary of these exciting developments:

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- lognormals and power laws can be awfully similar
- Random Multiplicative Growth leads to lognormal distributions
- leads to a power law tail eads to a power law tail
- With no minimum size but a distribution of lifetimes, the double Pareto distribution appears
- 🗞 Take-home message: Be careful out there...



References I

[1] R. Axtell. Zipf distribution of U.S. firm sizes. Science, 293(5536):1818–1820, 2001. pdf 7

[2] R. Gibrat. Les inégalités économiques. Librairie du Recueil Sirey, Paris, France, 1931.

[3] B. A. Huberman and L. A. Adamic. Evolutionary dynamics of the World Wide Web. Technical report, Xerox Palo Alto Research Center, 1999.

[4] B. A. Huberman and L. A. Adamic. The nature of markets in the World Wide Web. <u>Quarterly Journal of Economic Commerce</u>, 1:5–12, 2000. PoCS, Vol. 1 @pocsvox

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References



na a 24 of 26

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[6] M. Mitzenmacher. A brief history of generative models for power law and lognormal distributions. Internet Mathematics, 1:226–251, 2003. pdf

[7] E. W. Montroll and M. W. Shlesinger. On 1/f noise and other distributions with long tails. Proc. Natl. Acad. Sci., 79:3380–3383, 1982. pdf PoCS, Vol. 1 @pocsvox

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[8] E. W. Montroll and M. W. Shlesinger. Maximum entropy formalism, fractals, scaling phenomena, and 1/f noise: a tale of tails. J. Stat. Phys., 32:209–230, 1983.



