New Product Diffusion with Influentials and Imitators

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July 2005

Acknowledgements
We benefited from comments by David Bell, Albert Bemmaor, Xavier Drèze, Peter Fader, Donald Lehmann, Gary Lilien, Piero Manfredi, Paul Steffens, Stephen Tanny, Masataka Yamada, and audience members at the 2005 Marketing Science Conference. We also thank Peter Fader for providing the SoundScan CD sales data.

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Abstract

We model the diffusion of innovations in markets where the set of ultimate adopters consists of two segments: *influentials*, who are more in touch with new developments and affect another segment of pure *imitators*, whose own adoptions do not affect the influentials. This two-segment structure is consistent with several theories in sociology and diffusion research as well as many “viral” or “network” marketing strategies. There are four main results. (1) Diffusion in a mixture of influentials and imitators can exhibit a dip or “chasm” between the early and later parts of the diffusion curve. (2) The proportion of adoptions stemming from influentials need not decrease monotonically but may start increasing again. (3) Erroneously specifying a mixed-influence model to a mixture process where influentials act independently from each other can generate systematic changes in the parameter values reported in earlier research. (4) Empirical analysis of 33 different data series indicates that the two-segment model fits better than the standard mixed-influence, the Gamma/Shifted Gompertz, and the Weibull-Gamma models, especially in cases where a two-segment structure is likely to exist. Also, the two-segment model fits marginally better than the recent Karmeshu-Goswami mixed-influence model in which the coefficients of innovation and imitation vary across potential adopters in a continuous fashion.

Key words: Diffusion of innovations; social contagion; social structure.
1. Introduction

With the fragmentation of mass media, the deepening skepticism of consumers towards advertising and marketing, and the decreasing ability of salespeople to reach business customers, companies are under pressure to increase their marketing ROI through more astute targeting of resources. Meanwhile, at a broader societal level, technologies like Minitel and the Internet have boosted opportunities for information sharing among consumers, and a renewed appreciation has emerged for how social identification and status considerations pervade consumption (e.g., Ehrenreich 1990; Maffesoli 1996; Schor 1998). In response to these developments, marketers are rediscovering the importance of social contagion. This is especially so for new products which only the most involved and knowledgeable customers may be aware of at first, and for new products with considerable functional, financial, or social risk so mainstream customers are likely to seek information from peers.

The newly deployed “viral” and “network” marketing strategies often share two key assumptions: (1) some customers are more in touch with new developments, and (2) some (often but not always the same) customers’ adoptions and opinions have a disproportionate influence—direct or indirect—on others’ adoptions (e.g., Gladwell 2000; Moore 1995; Rosen 2000; Slywotzky and Shapiro 1993). Targeting those influential prospects who are more in touch with new developments and converting them into customers, the logic goes, allows marketers to benefit from a social multiplier effect on their marketing efforts. The two assumptions are quite reasonable, as they are consistent with several theories and a large body of empirical research (e.g., Katz and Lazarsfeld 1955; Keller and Berry 2003; Rogers 2003), and the social multiplier logic cannot be faulted either (e.g., Case et al. 1993; Valente et al. 2003). Yet, marketing science provides little or no additional theoretical or descriptive insight on how new products diffuse in
such markets. The reason is that the great majority of marketing diffusion models assume homogeneity rather than heterogeneity in the tendency to be in tune with new developments or the tendency to influence (or be influenced by) others, and often assume homogeneity in both. The present analysis addresses this gap between theory and emerging practice on the one hand, and marketing diffusion models on the other. Specifically, we model the aggregate-level diffusion path of a new product when the set of ultimate adopters is not homogenous but consists of two segments: influentials who are more in touch with new developments and who affect another segment of pure imitators whose own adoptions do not affect the influentials. We allow for the presence or absence of contagion among influentials.

While many diffusion models incorporate the dual drivers of independent decision making affected by being in touch with new developments and of imitation driven by others’ prior adoptions, they do so under the assumption that all potential adopters are ex ante affected equally by both factors. Taga and Isii (1959) in statistics, Mansfield (1961) and Williams (1972) in economics, Coleman (1964) in sociology, and Bass (1969) and Massy, Montgomery and Morrison (1970) in marketing, all advanced a model specifying the rate at which actors who have not adopted yet do so at time $t$ as $h(t) = p + qF(t)$, where $F(t)$ is the proportion of ultimate adopters that has already adopted, parameter $q$ captures social contagion, and parameter $p$ captures the time-invariant tendency to adopt early affected by consumer characteristics, the innovation’s appeal, and efforts of change agents.\(^1\) Since the proportion that adopts at time $t$ can be written as $f(t) = dF(t)/dt = h(t) [ 1 – F(t) ]$, one obtains:

$$f(t) = dF(t)/dt = [ p + qF(t) ] [ 1 – F(t) ]$$  [1]

The solution of this differential equation can be written as:

---

\(^1\) Following the convention in marketing, we refer to the rate at which non-adopters turn into adopters as the hazard rate and denote it as $h(t)$, even though the models we discuss are deterministic rather than probabilistic.
\[ F(t) = \frac{[1 - e^{g \cdot (p+q)t}]}{[1 + (q/p) e^{g \cdot (p+q)t}]} \]  

where \( g \) acts as a location parameter fixing the curve on the time axis (e.g., Mansfield 1961).

When \( t = 0 \) corresponds to the actual launch time such that \( F(0) = 0 \), then \( g = 0 \) and equation (2) reduces to the solution popular in marketing:

\[ F(t) = \frac{[1 - e^{(p+q)t}]}{[1 + (q/p) e^{(p+q)t}]} \]

The rate is influenced by both the intrinsic tendency to adopt \( (p) \) and social contagion \( (q) \) at all times except at \( t = 0 \) when \( qF(0) = 0 \). To reflect this dual influence, Mahajan and Peterson (1985) refer to the model as the mixed-influence model. Because the rate contains no contagion pressure at \( t = 0 \), those adopting at that time are sometimes referred to as innovators and contrasted against all others adopting later who are called imitators (e.g., Bass 1969). However, as several researchers have noted over the years, this terminology can be used only \textit{ex post} and the model does \textit{not} represent a diffusion process in an \textit{ex ante} mixture of two segments, the first adopting independently at a constant rate of \( p \) and the second segment driven only by social contagion and adopting at rate \( qF(t) \) (e.g., Bemmaor 1994; Jeuland 1981; Lekvall and Wahlbin 1973; Manfredi et al. 1998; Steffens and Murthy 1992; Tanny and Derzko 1988).

The objective of this study is, in the spirit of Bass (1969), Miller et al. (1993) and Moorthy (1993), to mathematically formalize theoretical arguments and behavioral research findings and to use this formalization to generate more refined theoretical insights on new product diffusion in a population of influentials and imitators. This is important as marketing practitioners increasingly deploy strategies assuming such a market structure and as marketing researchers increasingly incorporate social structure into their diffusion investigations (e.g., Bronnenberg and Mela 2004; Frenzen and Nakamoto 1993; Putsis et al. 1997; Van den Bulte and Lilien 2001).
Our results offer formalized insights into some current substantive and methodological research questions. First, diffusion in a mixture of influentials and imitators can exhibit a dip or “chasm” between the early and later parts of the diffusion curve. While this is a popular contention (e.g., Moore 1991), our model shows that it need not be necessary for firms to change their product to gain traction among later adopters and the adoption curve to swing up again. Like Steffens and Murthy (1992) and Karmeshu and Goswami (2001) but unlike Goldenberg et al. (2002), we obtain this result from a closed-form solution, and unlike those prior analyses, we show that a dip can occur even when influentials act independently from each other. Second, the proportion of adoptions stemming from influentials need not decrease monotonically but may start increasing again. The management implication is that, while it may make sense to shift the focus of one’s marketing efforts from influentials to imitators shortly after launch as shown by Mahajan and Muller (1998) using a two-period model, one may want to revert one’s focus back to influentials later in the process. Third, erroneously specifying a mixed-influence model to a two-segment process can generate the systematic changes in the parameter values over time reported in several studies. This analytical result is a specific formalization of Van den Bulte and Lilien’s (1997) more general but qualitative argument that unaccounted heterogeneity in \( p \) or \( q \) can generate changes in these parameters. Our result also complements Bemmaor and Lee’s (2002) simulation analysis since we consider heterogeneity in a process where genuine contagion exists rather than in a Gamma/Shifted Gompertz process without contagion.

We also perform an empirical analysis and assess the descriptive performance of the two-segment model compared to that of the mixed-influence model and of three diffusion models incorporating heterogeneity in the form of a continuous rather than a discrete mixture. Given the difficulty of unambiguously identifying causal processes from aggregate diffusion data
(Bemmaor 1994; Hernes 1976; Lekvall and Wahlbin 1973; Lilien et al. 1981; Van den Bulte and Stremersch 2004), the objective of this empirical analysis is not to conclusively demonstrate the validity of any model. Rather, it is to assess whether the differences between the discrete mixture and other models are sufficiently important to lead to differences in descriptive performance when applied to data of interest to marketing researchers. The two-segment model fits better than the mixed-influence, Gamma/Shifted Gompertz (Bemmaor 1994), and Weibull-Gamma models (Hardie et al. 1998; Massy et al. 1970; Narayanan 1992), especially in cases where a two-segment structure is likely (or even known) to exist, and fits slightly better than a recently advanced mixed-influence model where \( p \) and \( q \) vary across potential adopters in a continuous fashion (Karmeshu and Goswami 2001).

We proceed with first outlining our model setting, and within that context, discuss five theories and frameworks that suggest the existence of *ex ante* influential and imitators. Next, we develop a macro-level model of innovation diffusion in such a setting. Subsequently, we discuss how this model relates to the familiar mixed-influence model and to prior work on two-segment models by Jeuland (1981), Tanny and Derzko (1988) and Steffens and Murthy (1992). Finally, we report on the descriptive performance of the influential-imitator model compared to that of the mixed-influence and continuous-mixture models.

2. **Theories motivating a two-segment structure of influential and imitators**

The situation we model is the following. The set of eventual adopters has a constant size \( M \) and consists of two *a priori* different types of actors, influential and pure imitators. We use the subscripts 1 and 2 to denote each type, and the subscript \( m \) to denote the entire mixture population of adopters. We use \( \theta \) to denote the proportion of type 1 actors in the population of eventual adopters (0 ≤ \( \theta \) ≤ 1), and \( F(t) \) to denote the cumulative penetration. Finally, \( w \) denotes
the relative importance that imitators attach to influentials’ versus other imitators’ behavior (0 ≤ \( w \leq 1 \)). Each type’s adoption behavior is then captured by the following hazard functions:

\[
h_1(t) = p_1 + q_1 F_1(t) \quad [4]
\]

\[
h_2(t) = q_2[w F_1(t) + (1-w) F_2(t)] \quad [5]
\]

Note the double asymmetry in the influence process. First, type 1 may influence type 2, but the reverse is not true. Since, \textit{ex ante}, anyone of type 1 may influence anyone of type 2, we label the former \textit{influentials}, which is consistent with industry parlance (Keller and Berry 2003). Second, unlike type 1, type 2 is driven only by contagion, so we label them \textit{imitators}. Obviously, when \( \theta = 1 \) or \( q_2 = 0 \), everyone falls into segment 1 and the situation reduces to the mixed-influence model. Also, when \( \theta = 0 \), everyone falls into segment 2 and the situation reduces to the logistic model. When imitators put equal weight on all prior adoptions regardless of origin, we have \( h_2(t) = q_2 F_m(t) = q_2[\theta F_1(t) + (1-\theta) F_2(t)] \), i.e., \( w = \theta \) (as shown in Section 3).

The distinction between influentials and imitators is based on what drives their adoption behavior, not on whether they adopt early or late. Hence, the distinction is different from that of innovators vs. imitators in Bass (1969) and innovators vs. early adopters vs. early majority vs. late majority vs. laggards in Rogers (2003). Conceptually, causal drivers and time of adoption need not map one-to-one. Empirically, while those adopting early may act independently of others, and those adopting late may be subject to contagion, this is not always so: many early adoptions may be driven by contagion and the bulk of the late adoptions may stem from people not subject to social contagion (e.g., Becker 1970; Coleman et al. 1966).

Several theories and conceptual models suggest such a two-segment structure, though there is some disagreement on whether \( q_1 \) may be larger than zero. We first describe sociological arguments focusing on social character, social status, and social norms. Then, we turn to the two-
step flow hypothesis that focuses on interest in new developments, and finally to the chasm idea that focuses on enthusiasm for innovations versus risk aversion.

2.1. Social character

In his classic treatise on the changing nature of modern society, Riesman (1950) distinguished three types of social character: autonomous, inner-directed, and other-directed. The first two have in common the presence of clear-cut internalized goals, but differ as to whether these are consciously chosen (autonomous) or inculcated during youth by elders (inner-directed). Other-directed actors, in contrast, use their peers as their source of direction. The typology is in essence about conformity stemming from the need for approval and direction from others. Riesman worked on a broad social and cultural canvas and his typology is best used to refer to patterns of behavior found in a variety of specific contexts rather than to types of persons or personalities. Yet, his concepts have direct relevance for consumer behavior (e.g., Riesman 1950; Schor 1998). Some actors in some situations will exhibit autonomous or inner-directed adoption behavior independent from their peers (hence $q_1 = 0$), while others will exhibit other-directed behavior driven by social contagion from peers. Riesman did not narrowly specify who these peers are, and allowed them to be all of society (so $w = \theta$ being possible).

2.2. Status competition and maintenance

People buy and use products not only for functional purposes but also to construct a social identity, and to confirm the existence and support the reproduction of social status differences (Bourdieu 1984). A long-held idea in diffusion theory is that people seek to emulate the consumption behavior of their superiors and aspiration groups (e.g., Simmel 1971) and also quickly pick up innovations adopted by others of similar status if they fear that such adoptions might undo the present status ordering (Burt 1987). In short, actors tend to imitate the adoptions
of those of higher and similar social status.

Assuming one can divide the population in a high-status and a low-status group, status considerations suggest that both groups may exhibit contagion. Higher-status actors may imitate each other out of fear of falling behind ($q_1 \geq 0$), and lower-status actors will imitate to catch up. Whose adoptions the imitators act upon is not clear \textit{a priori}. One might argue that the only adoptions that matter are those by the high-status influentials to which they want to move closer ($w \rightarrow 1$). However, most authors follow Simmel and posit a finer-grained hierarchy with multiple strata (approximated imperfectly by a dichotomy) and a cascading pattern through the population where all prior adoptions contribute equally to social contagion ($w = \theta$). Finally, to the extent that status is maintained by adhering to social norms enforced among one’s direct peers of similar position, imitators should care mostly about fellow imitators ($w \rightarrow 0$).

2.3. Middle-status conformity

Like theories of status competition and maintenance, middle-status conformity theory is about one’s proper place in society. The main claim is that the relationship between status and conformity to norms—and hence susceptibility to social contagion—is an inverted U (e.g., Homans 1961; Philips and Zuckerman 2001). Since high-status actors feel confident in their social acceptance, they feel comfortable to deviate from conventional behavior and adopt appealing innovations independently from others. Low-status actors feel free to deviate from accepted practice and adopt innovations independently as well because they feel that this can not hurt their already low status. Middle-status actors, in contrast, feel insecure and strive to demonstrate their legitimacy by engaging in new practices only after they have been socially validated. So, middle-status conformity theory is consistent with the presence of two kinds of
actors, one adopting as a function of the innovation’s appeal irrespective of others’ actions \((q_1 = 0)\), and one adopting as a function of the legitimation stemming from prior adoptions.

The theory does not specify whose adoptions are being imitated \((w)\). Conformity pressures allow for both selective “status-based” imitation and non-selective “frequency-based” imitation (Haunschild and Miner 1997). Adoptions by high-status actors might legitimate the innovation in the eyes of the middle-status actors disproportionately, in which case the relation of \(w\) to \(\theta\) is unclear as the latter captures both high and low status. Conversely, imitators may care only about social acceptability among their middle-status peers, and hence care only about the latter’s adoptions \((w = 0)\). Finally, applications of neo-institutional theory to innovation adoption tend to posit that the legitimacy of an innovation is affected by the overall penetration rate \((w = \theta)\).

Note, higher status is often associated with higher economic resources and hence a higher ability to adopt innovations. This leads to the interesting prediction that only the adoptions at an intermediate stage of the overall diffusion process (made by middle-status actors) exhibit contagion (e.g., Cancian 1979), because the earliest adoptions will come from high-status actors and the latest from low-status actors, none of which are subject to contagion.

2.4. Two-step flow

The two-step flow hypothesis, originally proposed to explain unexpectedly weak mass media effects in presidential elections, posits that “ideas often flow from radio and print to the opinion leaders and from them to the less active sections of the population” (Lazarsfeld et al. 1944, p. 151; emphasis in original). So, in its original and starkest version, the two-step flow hypothesis posits two groups, one being affected only by mass media and the other being affected only by social contagion. What distinguishes the two groups is the level of interest in the subject matter and alertness to new developments rather than exposure to mass communications (Lazarsfeld et
al. 1944). Later studies in marketing have corroborated a strong relationship between opinion leadership and product interest and involvement (e.g., Coulter et al. 2002; Myers and Robertson 1972). Note, the two-step flow hypothesis does not impose that an opinion leader in one sphere (politics, fashion, computer games, etc.) also be a leader in another sphere, and several studies indeed document only moderate to little overlap in leadership across product categories (e.g., Katz and Lazarsfeld 1955; Merton 1949; Myers and Robertson 1972; Silk 1966). So, the relative size of the segments ($\theta$) may vary across innovations. While early studies focused on information flows from opinion leaders to less active members of the population, subsequent research has documented extensive information exchange among opinion leaders (e.g., Coulter et al. 2002; Katz and Lazarsfeld 1955). This would be consistent with $q_1 > 0$.

The two-step flow hypothesis emphasizes the flow of information. The contagion mechanism is one of information transfer increasing awareness of the product’s existence and decreasing its perceived risk, not of normative legitimation or status competition. Of the five theories we consider, this is perhaps the most familiar to marketers and the most flexible. For low-risk innovations, for instance, the fraction of pure imitators in need of guidance can be quite small, and $\theta$ quite large. Who is being imitated is not clearly specified, and $w$ may range from 0 to 1. The two-step flow idea emphasizes that mass media influence on the less-active segment operates through opinion leaders who are the only ones to take an active interest in information available in the media. It does so without constraining the social influence exerted on the less-active segment to come only from opinion leaders, and allows for a cascading or rolling pattern through the population where all prior adoptions contribute to social contagion (e.g., Katz 1957; Merton 1949). This suggests $w \approx \theta$. However, it is quite possible that opinion leaders are more influential, suggesting that—in the extreme case—they may be the only ones being imitated ($w = \ldots$)
1). Conversely, it is also quite possible that imitators consider fellow imitators to be more representative and hence valuable as information sources, suggesting low values of $w$.

**2.5. High-technology adoption chasm**

In Moore’s (1991) chasm framework for technology products, the so-called early market consists of “technology enthusiasts” and “visionaries” who are quick to appreciate the nature and benefits of the innovation, whereas the “mainstream” market consists of more risk-averse decision makers and firms who fear being stuck with a technology that is not user friendly, poorly supported, or at risk of losing a standards war. Whereas the mainstream market can be represented as responding only to the size of the installed base, i.e., prior adoptions (Mahajan and Muller 1998), Moore is unclear about the process among “technology enthusiasts” and “visionaries”. Whereas his textual discussions suggest that they act independently ($q_1 = 0$), his stylized graph of the bell-shaped adoption curve with a chasm is mathematically inconsistent with a constant-hazard process in the early stages of diffusion and requires $q_1 > 0$.

Moore does not clearly specify whose adoptions are being imitated ($w$). On the one hand, one might argue that the legitimacy of a new technology is affected by the penetration rate in the overall population, i.e., the total installed base regardless of who adopted ($w = 0$). On the other hand, Moore emphasizes that product versions appealing to technology enthusiasts and visionaries need not appeal to the mainstream market, which implies that mainstream customers discount adoptions by technology enthusiasts and visionaries and care only about adoptions by other mainstream customers ($w = 0$).²

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² Moore himself is far from clear on the issue when discussing the relationship between “visionaries” in the early market and “pragmatists,” i.e., the early adopters among the members of the mainstream market. At one point, he admonishes the reader to “do whatever it takes to make [visionaries] satisfied customers so that they can serve as good references for the pragmatists” but on the very next page he writes that “pragmatists think visionaries are dangerous. As a result, visionaries, with their highly innovative … projects do not make good references for pragmatists” (Moore 1995, pp. 18-19).
2.6. Conclusion

At least five different theoretical frameworks imply modeling innovation diffusion using a two-segment structure consisting of influentials and imitators (Table 1). Two theories suggest that influentials adopt independently, implying $q_1 = 0$, but the other three suggest that influentials may exhibit contagion amongst themselves.\(^3\) The theories vary in their causal mechanisms and, consequently, in what kind of actors belongs to each segment and who the imitators imitate ($w$).

<table>
<thead>
<tr>
<th>Framework</th>
<th>Influentials</th>
<th>Imitators</th>
<th>Reason to imitate</th>
<th>Who gets imitated $^a$</th>
</tr>
</thead>
<tbody>
<tr>
<td>Social character</td>
<td>Autonomous and inner-directed; $q_1 = 0$</td>
<td>Other-directed</td>
<td>Looking for approval and direction</td>
<td>- Not specified, possibly all adopters ($w = \theta$)</td>
</tr>
<tr>
<td>Status competition and maintenance</td>
<td>High status; $q_1 \geq 0$</td>
<td>Low status</td>
<td>Gaining or maintaining status</td>
<td>- All adopters ($w = 0$)</td>
</tr>
<tr>
<td></td>
<td>Middle-status conformity</td>
<td>High and low status; $q_1 = 0$</td>
<td>Middle status</td>
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</tr>
<tr>
<td>Two-step flow</td>
<td>Active and involved (opinion leaders); $q_1 \geq 0$</td>
<td>Not active or involved</td>
<td>Transferring information</td>
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</tr>
<tr>
<td>Technology chasm</td>
<td>Technology enthusiasts and visionaries; $q_1 \geq 0$</td>
<td>Mainstream customers</td>
<td>Reducing risk</td>
<td>- All adopters ($w = 0$)</td>
</tr>
</tbody>
</table>

$^a$ Parameter $w$ denotes how much the social contagion affecting the imitators stems from the influentials ($w$) rather than fellow imitators ($1-w$). Parameter $\theta$ is the fraction of ultimate adopters belonging to segment 1 (influentials).

The theories also suggest that the relative size of the segments ($\theta$) can vary from innovation to innovation. It may be quite low for very non-mainstream products that only a very small pocket of “bleeding edge” customers find attractive but that in spite of the latter’s enthusiasm take a long time to diffuse, resulting in an adoption curve with a long left tail. Conversely, for products with low functional or financial risk and with little implications for social status, like

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\(^3\) Independent decision making among influentials is also consistent with Midgley and Dowling (1978) who define innovativeness as “the degree to which an individual makes innovation decisions independently of the communicated experience of others” (p. 235). So our distinction between independent influentials (with $q_1 = 0$) and imitators is the same as their dichotomy between “innate innovators” and “innate noninnovators” (p. 237).
marginally novel drugs or CDs and movies with already famous performers, most adopters may feel little need for information or legitimation from peers. This implies a high $\theta$, a low $q_1$, and an exponential-like diffusion process (e.g., Moe and Fader 2001; Van den Bulte and Lilien 2001). A high $\theta$ is also expected when the large majority of adopters behave according to the mixed-influence model, with consumer appliances being a likely candidate.

3. Two-segment mixture models

We seek closed-form solutions in the time domain for an innovation’s diffusion path when the set of eventual adopters, which has a constant size $M$, consists of two a priori different types of actors adopting according to equations (4) and (5). The overall cumulative penetration is simply the average of both types’ cumulative penetration weighted by their constant population weights (e.g., Cox 1959):

$$F_m(t) = \theta F_1(t) + (1-\theta) F_2(t) \quad [6]$$

Similarly, the fraction of the population adopting at time $t$ is:

$$f_m(t) = \theta f_1(t) + (1-\theta) f_2(t) \quad [7]$$

In contrast, the population hazard function is not an average of the two hazards weighted by each segment’s constant population weights, but is given by:

$$h_m(t) = f_m(t) / [1-F_m(t)]$$

$$= \left[ \theta f_1(t) + (1-\theta) f_2(t) \right] / [1-F_m(t)]$$

$$= \pi(t) h_1(t) + [1-\pi(t)] h_2(t) \quad [8]$$

where $f_i(t) = h_i(t) [1-F_i(t)]$ and $\pi(t)$ is the proportion of actors not having adopted yet at time $t$ that belong to type 1:

$$\pi(t) = \frac{\theta}{1-F_m(t)} \frac{1-F_i(t)}{1-F_m(t)} \quad [9]$$
Finally, the proportion of adoptions taking place at time $t$ that is made by actors of type 1 is:

$$
\phi(t) = \frac{\theta f_1(t)}{f_m(t)} \quad [10]
$$

### 3.1. Influential-imitator mixture model (IIM), with $q_1 > 0$

Having defined the key functions, and having made the behavioral assumptions in the hazard functions (eqs. 4 and 5), we now develop the influential-imitator mixture model (IIM). The process among the influentials is the well-known mixed-influence model. When $F_1(0) = 0$, the cumulative penetration function and instantaneous adoption function for influentials are:

$$
F_1(t) = (1 - e^{-(p_1 + q_1)t})/(1 + \frac{q_1}{p_1} e^{-(p_1 + q_1)t}) \quad [11]
$$

$$
f_1(t) = (p_1 (1 + \frac{q_1}{p_1})^2 e^{-(p_1 + q_1)t})/(1 + \frac{q_1}{p_1} e^{-(p_1 + q_1)t})^2 \quad [12]
$$

The diffusion path among imitators, in contrast, does not follow any standard diffusion model, as it is driven by the prior adoptions of both influentials and other imitators. As shown in Appendix A1, when $F_2(0) = 0$, the cumulative penetration function for imitators in IIM is:

$$
F_2(t) = 1 + \frac{p_1 w - q_1 (1 - w)}{q_1 (1 - w)H_1 + e^{q_1 t} \left( \frac{p_1 + q_1 e^{-(p_1 + q_1)t}}{p_1 + q_1} \right)^\frac{1}{n} (q_1 (1 - w) - p_1 w - q_1 (1 - w)H_2)} \quad [13]
$$

where

$$
H_1 = F_1(1, \frac{w_2}{q_1} ; 1 + \frac{w_2}{q_1} ; \frac{p_1}{p_1 + q_1} ; \frac{p_1}{p_1 + q_1} e^{-(p_1 + q_1)t}) \quad , \quad H_2 = F_1(1, \frac{w_2}{q_1} ; 1 + \frac{w_2}{q_1} ; \frac{q_1}{p_1 + q_1} ; \frac{p_1}{p_1 + q_1}) \quad , \quad H_1 \quad and \quad H_2
$$

$$_2F_1(1,b;c;k)$$ is the Gaussian hypergeometric function:

$$
_2F_1(1,b;c;k) = \sum_{n=0}^{\infty} \frac{\Gamma(b+n)\Gamma(c)}{\Gamma(b)\Gamma(c+n)} k^n \quad [14]
$$

This hypergeometric series is convergent for arbitrary $b, c$ if $|k| < 1$; and for $k = \pm 1$ if $c > 1 + b$.

This implies that the closed-form solution in equation (13) is well-defined as long as $q_1 > 0$.\footnote{While the Gaussian hypergeometric functions $_2F_1(1,b;c;k)$ can be simplified to incomplete beta functions, we do not perform this simplification as it requires the overly restrictive condition that $p_1 w/(1-w) > q_1$.}

Once $F_1(t)$ and $F_2(t)$ are known, one can obtain the instantaneous adoption function $f_2(t)$ by
substituting equations (11) and (13) into:

\[ f_2(t) = q_2 \left[ wF_1(t) + (1-w) F_2(t) \right] [1- F_2(t)] \]  \[15\]

With solutions for \( F_1(t), f_1(t), F_2(t) \) and \( f_2(t) \) available, one can enter those into equations (6) through (10) to obtain closed-form solutions for the population-level functions.\(^5\) In Figure 1, we plot the function \( f_m(t) \) and its two components \( \theta f_1(t) \) and \((1-\theta)f_2(t)\) for four sets of parameter values chosen to illustrate various types of diffusion behavior possible in this model:

Case (a): \( p_1 = 0.05; q_1 = 0.1; q_2 = 0.2; \theta = 0.15; w = 0.20 \);
Case (b): \( p_1 = 0.01; q_1 = 0.5; q_2 = 0.2; \theta = 0.15; w = 0.01 \);
Case (c): \( p_1 = 0.05; q_1 = 0.5; q_2 = 0.2; \theta = 0.30; w = 0.30 \);
Case (d): \( p_1 = 0.01; q_1 = 0.1; q_2 = 0.2; \theta = 0.15; w = 0.001 \).

Diffusion process (a) exhibits a bell-shaped adoption curve \( f_m(t) \) that is unimodal and close to symmetric around its peak. This is the pattern commonly associated with mixed-influence model. Diffusion process (b) is bimodal and exhibits a marked dip because adoptions by influentials are already well past their peak by the time the imitators start adoption in numbers (the delay being caused by the low \( w \) value). This is the much-debated “chasm” pattern. Diffusion processes (c) and (d), finally, are again unimodal but exhibit a clear skew to the right or left, which the mixed-influence cannot account for very well (e.g., Bemmaor and Lee 2002).\(^6\) Note that in all four cases, \( f_1(t) \) reaches zero before \( f_2(t) \) does, so the commonly expected association between being

---

\(^5\) Even though our closed-form solution for \( F_2(t) \) in the IIM looks quite different from the solution presented by Steffens and Murthy (1992), theirs is actually nested in ours. After imposing the constraint \( w = \theta \), reparameterizing the Steffens-Murthy solution in terms of \( m, \theta, p_1, q_1 \), and \( q_2 \), correcting for a (most likely typographic) error in their solution, and performing additional derivations, one can show that our closed-form solution for \( F_2(t) \) in the IIM, and hence \( f_m(t) \), is identical to theirs. One difference, though, is that their solution requires \( q_1 > q_2 \theta \) (or \( q_1 > q_2 w \)) for a series expansion term in their solution to converge, whereas the solution in eq. (13) only requires \( q_1 > 0 \).

\(^6\) All four patterns for the total number of adoptions shown in Figure 1 have been documented in prior research. Pattern (a) is probably the most commonly reported in the marketing literature. Steffens and Murthy (1992) and Karmeshu and Goswami (2001) report data series exhibiting the bimodal pattern (b). Dixon (1980) reports the presence of long right tails, i.e., pattern (c), in many of the data he analyzed. Van den Bulte and Lilien (1997) report several data series exhibiting long left tails, i.e., pattern (d).
an imitator and being a late adopter holds. Also note that, as one would intuit, low values of $w$ cause the diffusion among imitators to be delayed and $f_2(t)$ to shift to the right. We now turn to the case where $q_1 = 0$, and study it in some more detail using the functions $h_m(t)$, $\pi(t)$, and $\phi(t)$.

**Figure 1: Adoption functions for four IIM diffusion processes**

3.2. Pure-type mixture model (PTM), with $q_1 = 0$

When influentials adopt independently and $q_1 = 0$, the situation is that of a pure-type mixture (PTM) of pure independents and pure imitators. The process among the independents is the well-known constant-hazard exponential process. When $F_1(0) = 0$, we have:

\[
F_1(t) = 1 - e^{p_1 t} \tag{16}
\]

\[
f_1(t) = p_1 e^{p_1 t} \tag{17}
\]

As shown in Appendix A2, when $F_2(0) = 0$, the cumulative penetration function for imitators in the PTM is:
\[ F_2(t) = 1 + \frac{\exp(-q_2 t - \frac{q_2}{p_1} we^{-p_1 t})}{\frac{q_2}{p_1} (1-w)} \left( \frac{q_2}{p_1} (\Gamma(\frac{q_2}{p_1}, \frac{q_2}{p_1} we^{-p_1 t}) - \Gamma(\frac{q_2}{p_1}, \frac{q_2}{p_1} w)) - \exp(-\frac{q_2}{p_1} w) \right) \]  

where \( \Gamma(\eta, k) \) is the “upper” incomplete gamma function:

\[
\Gamma(\eta, k) = \int_{k}^{\infty} v^{\eta-1} e^{-v} dv
\]

The instantaneous adoption function \( f_2(t) \) is obtained by substituting equations (16) and (18) into (15). With solutions for \( F_1(t), f_1(t), F_2(t) \) and \( f_2(t) \) available, one can enter those into equations (6) through (10) to obtain closed-form solutions for the population-level functions. In Figure 2, we plot the functions \( f_m(t), h_m(t), \pi(t), \phi(t) \) for three sets of parameter values chosen to illustrate various types of diffusion behavior possible in this model:

Case (a): \( p_1 = .15, q_2 = .50, \theta = .25, w = .25 \);

Case (b): \( p_1 = .25, q_2 = .40, \theta = .15, w = .01 \);

Case (c): \( p_1 = .15, q_2 = .65, \theta = .60, w = .05 \).

Diffusion process (a) exhibits the common unimodal, symmetric-around-the-peak adoption curve \( f_m(t) \) well captured by the mixed-influence model. More interesting is that the hazard function is not monotonic as in the mixed-influence model. Rather, it is roughly bell-shaped and seems to converge to a value in between the minimum and the maximum. Here is why. The very earliest adopters consist of independents and the population hazard equals \( \theta p_1 = .0375 \) at first. As more and more imitators adopt with hazard \( q_2 F_m(t) \), the population hazard increases. Once \( q_2 F_m(t) > p_1 \), which can happen quickly when \( q_2 \) is markedly larger than \( p_1 \), the set of imitators not having adopted yet will start depleting faster than the set of independents not having adopted.

---

7 Of the three shapes of adoption curve in Figure 2, pattern (a) is probably the most commonly reported in the diffusion literature. The other two shapes have not been documented as extensively, but do occur in previously analyzed data. For instance, the sales curve of several music CDs studied by Moe and Fader (2001) exhibit pattern (b) or (c), and the classic Medical Innovation data analyzed by Coleman et al. (1966) also exhibit pattern (c).
Figure 2: Plots of functions characterizing three PTM diffusion processes

(a) $p_1 = .15, q_2 = .5, \theta = .25, w = .25$

(b) $p_1 = .25, q_2 = .4, \theta = .15, w = .01$

(c) $p_1 = .15, q_2 = .65, \theta = .6, w = .05$
yet. As a result, the laggards remaining to adopt consist increasingly of independents—as indicated by the function $\pi(t)$ reaching a minimum around $t = 5$ and then increasing to 1—and the population hazard converges back to an asymptote of $p_1 = .15$. This pattern of relative speed of depletion also explains the non-monotonic pattern in $\phi(t)$, the proportion of adoptions taking place at time $t$ stemming from independents. Note that in this diffusion process, independents make up the bulk not only of the early adopters, but also of the very late adopters. Importantly, the point at which $\phi(t)$ starts increasing and independents start gaining rather than losing importance ($t = 7.3$) occurs when the process is still far from complete and the remaining market potential is still quite sizable (37 % since $F_m(t) = .63$ at $t = 7.3$).

Diffusion process (b) differs in several respects from process (a). First, the adoption curve $f(t)$ does not have a smooth bell shape but exhibits a clear dip early on. This is easily explained. The independents adopt rapidly because $p_1 = .25$ is rather high. However, imitators’ reaction to those independent adoptions is very muted because they imitate mostly fellow imitators ($w = .01$). As a result, the adoptions by independents show an exponential decline which is not immediately compensated by the imitators’ slowly developing adoptions, resulting in an early dip in the population curve. Note that independents account for the bulk of the adoptions only early in the diffusion process, as $\phi(t)$ declines steeply to close to zero. So, while the adoption curve does not fit the standard model, we do have the commonly expected association between being an imitator and being a late adopter.

Diffusion process (c) is mostly an exponential process commonly observed for fast moving consumer goods, CDs and films, but with a marked boost after the early periods. What is happening is that most adopters are independents ($\theta = 60\%$), so the majority of adoptions follow an exponential decline. However, there is also a sizable segment of imitators that are very
sensitive to social contagion \((q_2 = .65)\), but mostly from fellow imitators rather than independents \((w = .05)\). As a result, the imitators are slow to adopt at first, but once the snowball starts rolling, tend to adopt in a very short time. This is reflected in the shape of \(\phi(t)\): the proportion of adoptions accounted for by independents tends to be close to 100%, except for a relatively narrow time window during which it first declines and then increases again. The contrast between process (a) and (c) is informative: They have similar \(p_1\) and \(q_2\) values, and the composition of both adopters \(\phi(t)\) and remaining non-adopters \(\pi(t)\) tend to evolve similarly, as do their respective population hazard functions \(h(t)\). Yet, because of the different segment sizes \(\theta\) and contagion weights \(w\) in the two processes, the resulting adoption curves are quite different.

### 3.3. Some special cases of theoretical interest

Our review of prior theories and frameworks indicates that three cases of the IIM and PTM are of special theoretical interest. The first is where imitators imitate only influentials \((w = 1)\) such that \(h_2(t) = q_2 F_1(t)\). The second is where imitators imitate only other imitators \((w = 0)\) such that \(h_2(t) = q_2 F_2(t)\). The third is where imitators randomly mix with both independents and imitators such that \(w = \theta\) and \(h_2(t) = q_2 F_m(t)\). In the first and third case, \(F_2(t)\) and \(f_2(t)\) are easily derived by imposing \(w = 1\) and \(w = \theta\), respectively, in equations (13), (15) and (18). In the second case, the process among imitators is only a function of prior adoptions by other imitators and is simply the well-known logistic process.\(^8\) In all three special cases, the population-level functions \(F_m(t), f_m(t), h_m(t), \pi(t),\) and \(\phi(t)\) are readily obtained once \(F_2(t)\) is known.

\(^8\) Note, when \(w = 0\) or \(\theta = 0\), the process among imitators cannot get started within the model. As is well known, the closed-form solution for the logistic requires that \(F_2(0) > 0\). Hence, while the cases with \(w = 0\) or \(\theta = 0\) are conceptually nested within IIM or PTM, their closed-form solutions are not as they make different assumptions about the initial conditions.
4. Relation to prior diffusion models

4.1. Mixed-influence model vs. pure-type mixture model

As the closed-form solutions and the plots in Figure 2 indicate, the mixed-influence model (MIM) does not capture diffusion processes in a discrete mixture of pure independents and pure imitators, regardless of the relative contagion influence of adoptions by independents vs. imitators. The only two exceptions to this are the case where \( p_1 = 0 \) or \( \theta = 0 \) and both models collapse to the logistic model, and the case where \( q_2 = 0 \) or \( \theta = 1 \) and both models collapse to the exponential model.

Our analysis allows one to assess more rigorously the argument (e.g., Mahajan et al. 1993) that rewriting the standard differential equation for the mixed influence model (eq. 1) into:

\[
f(t) = p \left[ 1 - F(t) \right] + qF(t) \left[ 1 - F(t) \right]
\]

allows one to interpret the term \( p \left[ 1 - F(t) \right] \) as the number adoptions made by people adopting with hazard \( p \) and the term \( qF(t) \left[ 1 - F(t) \right] \) as the number of adoptions made by people adopting with hazard \( qF(t) \). While the manipulation of the equation is evidently correct, the interpretation is not. The main reason is that, in each term, the fraction of actors not having adopted yet, \( 1 - F(t) \), refers to the total population, rather than to the fractions in each of the segments, \( 1 - F_1(t) \) and \( 1 - F_2(t) \). In addition, the sizes of each segment are ignored. The correct expression for a mixture, is:

\[
f_m(t) = \theta f_1(t) + (1-\theta) f_2(t)
\]

\[
= \theta h_1(t) [1-F_1(t)] + (1-\theta) h_2(t) [1-F_2(t)]
\]

\[
= \theta p_1 \left[ 1 - F_1(t) \right] + (1-\theta) q_2 [wF_1(t) + (1-w) F_2(t)] \left[ 1 - F_2(t) \right]
\]

When imitators randomly mix with independents and imitators and are equally affected by both, then \( w = \theta \) and the equation simplifies to:

\[
f_m(t) = \theta p_1 \left[ 1 - F_1(t) \right] + (1-\theta) q_2 F_m(t) \left[ 1 - F_2(t) \right]
\]
Even if $p = \theta p_1$, $q = (1-\theta)q_2$, and one omits the $m$-subscript from the population-level $f_m(t)$ and $F_m(t)$, the mixture equation (21) is different from the mixed-influence equation (19).

Within a homogeneous population with mixed influence, one can only interpret the relative size of the two terms $p[1-F(t)]$ and $qF(t)[1-F(t)]$ as reflecting the relative influence of time-invariant elements ($p$) versus social contagion ($qF(t)$) on the adoptions at time $t$, keeping in mind that each and every adoption is influenced by both $p$ and $qF(t)$ for any $t > 0$. For instance, the ratio $p/(p+qF(t))$ can be used as a measure of the relative strength of time-invariant elements at time $t$ (Lekvall and Wahlbin 1973), as can the decomposition presented by Daley (1967) and Mahajan, Muller and Srivastava (1990), but neither can be interpreted as the fraction of all adoptions at time $t$ stemming from pure-type actors adopting \emph{a priori} with hazard $p$.

Another common belief about the mixed-influence model that is inconsistent with its mathematical structure is that “the importance of innovators will be greater at first but will diminish monotonically with time,” where innovators are defined as those who “are not influenced in the timing of their initial purchase by the number of people who have already bought the product” (Bass 1969, p. 217). In a homogenous population where everyone behaves according to the hazard rate $p + qF(t)$, the only actors with hazard $p$ are those adopting at $t = 0$ when $F(0) = 0$. Anyone adopting afterwards is influenced by prior adoptions. Hence, in the mixed-influence model, the proportion of adoptions occurring at time $t$ that are unaffected by social contagion follows a step function with value 1 at $t = 0$ and value 0 for any $t > 0$.

Conversely, in a true discrete mixture, the proportion of independents adopting with a constant hazard, i.e., function $\phi(t)$, need not diminish monotonically over time, as shown in Figure 2.
4.2. Consequence of imposing a mixed-influence structure on a pure-type mixture process

From comparing equations (19) and (21) one may get the impression that a diffusion process in a discrete mixture with \( h_1(t) = p_1 \) and \( h_2(t) = q_2 F_m(t) \) could be approximated quite well by a mixed-influence model with \( h(t) = p + qF(t) \), even if they are not identical. However, the adoption functions \( f_m(t) \) and hazard functions \( h_m(t) \) suggest some potentially important deviations. More insight comes from re-writing the expression for \( f_m(t) \) in eq. (21) into a form similar to that for \( f(t) \) in the mixed-influenced model (following Manfredi et al. 1998):

\[
\begin{align*}
f_m(t) &= \theta p_1 [ 1 - F_1(t) ] + (1- \theta) q_2 F_m(t) [ 1 - F_2(t) ] \\
&= \left[ \theta p_1 \frac{1-F_1(t)}{1-F_m(t)} + (1- \theta) q_2 F_m(t) \frac{1-F_2(t)}{1-F_m(t)} \right] [ 1 - F_m(t) ] \\
&= \left[ p(t) + q(t) F_m(t) \right] [ 1 - F_m(t) ] \tag{22}
\end{align*}
\]

where

\[
\begin{align*}
p(t) &= \theta \frac{1-F_1(t)}{1-F_m(t)} p_1 = \pi(t) p_1 \tag{23} \\
q(t) &= (1-\theta) \frac{1-F_2(t)}{1-F_m(t)} q_2 = [1-\pi(t)] q_2 \tag{24}
\end{align*}
\]

Deleting the \( m \) subscript from equation (22) to reflect one’s ignoring that the population consists of a mixture results in:

\[
f(t) = \left[ p(t) + q(t) F(t) \right] [ 1 - F(t) ] \tag{25}
\]

So, one is able to re-write the pure-type mixture model with \( w = \theta \) into an expression akin to the mixed-influence model, but with both hazard rate parameters varying systematically over time. More specifically, \( p(t) \) changes in exactly the same way as \( \pi(t) \), the proportion of actors not having adopted yet by time \( t \) that belong to the segment of independents. At \( t = 0 \), \( \pi(t) = \theta \) and \( p(t) = \theta p_1 \). Since at the very beginning adoption tends to be more prevalent among independents
than among imitators, the number of independents who have not adopted yet gets depleted faster than the number of imitators who have not. Consequently, $\pi(t)$ and $p(t)$ decline at first. However, when $q_2 >> p_1$, the relative speed of adoption between the two segments quickly reverses and the set of actors who have not adopted yet tends to become increasingly dominated by independents. As a result, $\pi(t)$ and $p(t)$ increase over most of the time window. The reverse pattern takes place for $q(t) = [1 - \pi(t)] q_2$. It starts at $(1-\theta)q_2$, increases for a very short period, but starts decreasing very soon. Note, when $\theta \approx 0$ or $\theta \approx 1$, then $\pi(t)$ will not vary much and neither will $p(t)$ or $q(t)$.

In short, specifying a mixed-influence model with $h(t) = p + qF(t)$ when the true data generating process is that of a discrete mixture with $h_1(t) = p_1$ and $h_2(t) = q_2F_m(t)$ where $q_2 >> p_1$ will yield increasing values of $p$ and decreasing values of $q$ (except for the first very few periods). This is consistent with the pattern in mixed-influence model estimates described in prior research. Though Van den Bulte and Lilien (1997) focused their analysis on ill-conditioning in the absence of model misspecification, they recognized that unobserved heterogeneity in $p$ and $q$ forms an alternative explanation for the systematic changes they observed in empirical applications. Our results formalize their argument for the case of two segments where one segment has $p = 0$ and the other has $q = 0$.

4.3. Relation to other two-segment models

Figure 3 shows how our models relate to a few other models, including two earlier two-segment models. Tanny and Derzko (1988) used a discrete mixture with $h_1(t) = p_1$ and $h_2(t) = p_2 + q_2F_m(t)$. Steffens and Murthy (1992) used a discrete mixture with $h_1(t) = p_1 + q_1F_1(t)$ and $h_2(t) = q_2F_m(t)$. So, as shown in Figure 3, both these models conceptually nest both the mixed-influence model and PTM3 with $w = 0$. The diagram also shows that, like the mixed-influence model, the pure-type mixture models have both the exponential and logistic models nested in
them, with the exception that PTM1 with \( w = 1 \) does not nest the logistic because if \( h_2(t) = q_2 F_1(t) \) and either \( \theta = 0 \) or \( p_1 = 0 \), then \( h_2(t) \) is undefined. Note, only the PTMs feature two “pure types,” i.e., independents and imitators without any mixed influence.

**Figure 3. Relations among the IIM and PTM models, the Steffens-Murthy and Tanny-Derzko models, and the mixed-influence, exponential, and logistic models**

A model receiving an arrow is conceptually nested in the model where the arrow originates. For instance, the general PTM with \( w = 1 \) generates PTM1 and the PTM1 with \( q_2 = 0 \) generates the exponential.

As shown in Appendix A3, the solution for \( F_2(t) \) in PTM3 is consistent with Jeuland’s (1981, p. 14) earlier work. The differences are that he did not specify \( h_1(t) \) but kept it general and that his partial solution still contained unknown integrals. In contrast, we specify the process among
independents and solve the equations using incomplete gamma functions, making parameter estimation and empirical analysis possible.

5. Empirical analysis

To what extent does the two-segment IIM, consistent with several theoretical frameworks, agree with empirical diffusion patterns? And how well does it do compared to the mixed-influence model and other, more flexible, models? We provide insights on those issues through an empirical analysis of 33 data series.

5.1. Data

One must use an informative variety of data sets if one is to draw sound conclusion on model performance. We therefore analyze four sets of data. The first consists of a single series on the diffusion of the broad-spectrum antibiotic tetracycline among 125 Midwestern physicians over a period of 17 months in the mid-1950s. This series comes from the classic Medical Innovation study (Coleman et al. 1966). It warrants special attention because it is commonly accepted as an instance of diffusion in a mixture of independents and imitators (e.g., Jeuland 1981; Lekvall and Wahlbin 1973; Rogers 2003).

The second set of data series consists of 19 music CDs, also a category where a two-segment structure is a priori likely to exist. Some customers are dedicated fans buying products by their favorite performers almost unconditionally (so \( q_1 = 0 \) is quite possible), while others end up buying the CD only after it has become popular and a must-buy (Farrell 1998; Yamada and Kato 2002). We use the weekly U.S. sales data analyzed previously by Moe and Fader (2001).\(^9\) Since people are very unlikely to buy two identical CDs for themselves or to replace an older copy, the sales data are unlikely to be contaminated by multiple or repeat purchases and can be treated as

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\(^9\) The full set consists of 20 data series, but we deleted one that still had not reached the time of peak sales.
Figure 4 shows the data of four CDs each illustrating one typical path: a rather smooth decline for Blind Mellon, an early dip followed by a recycle for AdamAnt, a slowly developing “sleeper” pattern for Everclear, and a bell shape for Dink.

**Figure 4. Weekly sales (adoption) data for four CDs**

The third set of data consists of five series of high-technology products, for which a two-segment structure with \( q_1 > 0 \) is quite possible (e.g., Moore 1991). The first three series consists of adoptions of CT scanners, ultrasound and mammography equipment among hospitals of all sizes (Van den Bulte and Lilien 1997). The fourth series consists of the penetration between 1979 and 1993 of CT scanners among hospitals with 50 to 99 beds. Controlling for size may be important, as larger hospitals have larger budgets and more highly skilled staff, and these differences may mask genuine contagion processes (e.g., Davies 1979). The fifth series consists of the penetration of personal computers among US households. The series covers the years
1981-1996, but to avoid left-censoring artifacts we impose 1975 as the actual launch year. The first three series are roughly bell-shaped, the latter two series show two “bells” separated by a dip or “chasm”.

The final set is a miscellaneous mix of 8 data series analyzed previously by Van den Bulte and Lilien (1997) and Bemmaor and Lee (2002) (these studies also included the tetracycline and three of the high-tech series). There is no compelling \textit{a priori} reason to expect a mixture of independents and imitators to be able to account better for those diffusion data than traditional models, and several innovations need not have diffused through contagion at all (Griliches 1962; Van den Bulte and Stremersch 2004). The adoption curves all have a very pronounced bell shape, with several showing skew that the MIM cannot account for (Bemmaor and Lee 2002).

\section*{5.2. Parameter estimates}

Though we have closed-form solutions for both IIM and PTM, the solution to the IIM involves Gaussian hypergeometric functions the estimation of which is very troublesome.\footnote{Nonlinear regression in R and Mathematica either did not converge at standard convergence criteria or enabled us to obtain point estimates but not standard errors. We experienced these problems even with simulated data, which rules out model misspecification as an explanation for these difficulties. Maximum likelihood estimation is known to be troublesome as well, even when the parameters of interest enter the function linearly rather than non-linearly as in the IIM (e.g., Fader et al. 2005).} We therefore estimate the IIM not using the standard Srinivasan-Mason (1986) estimation approach but through direct integration, that is, we compute non-linear least squares estimates at the same time as we numerically solve the following differential equation\footnote{This can be done conveniently, e.g., using the model procedure in SAS or the odesolve package in R.}:  

\[
dX(t)/dt = M \left[ \theta f_1(t) + (1-\theta)f_2(t) \right] + \sigma(t) \\
= M \left[ \theta f_1(t) + (1-\theta) q_2 \left\{ wF_1(t)+(1-w) \frac{X(t)/M - \theta F_1(t)}{1-\theta} \right\} \left\{ 1- \frac{X(t)/M - \theta F_1(t)}{1-\theta} \right\} \right] + \sigma(t)
\]

where \(X(t)\) is the cumulative number of adopters observed at time \(t\), \(f_1(t)\) and \(F_1(t)\) are the closed-form solutions to the adoption and penetration functions of the MIM, and \(f_2(t)\) is expressed as in
eq. (15), but with \( \frac{X(t)}{M - \theta F_1(t)} \) replacing \( F_2(t) \). The latter is based on \( X(t) = MF_m(t) \) (absent error) and \( F_m(t) = \theta F_1(t) + (1-\theta)F_2(t) \). We allow the error term \( \epsilon(t) \sim N(0, \sigma^2) \) to exhibit serial correlation up to order 2 when the time series contains more than 20 observations or the Durbin-Watson statistic falls outside the 1.5-2.5 range. We impose that hazard parameters \( p_1, q_1, \) and \( q_2 \) be non-negative (\( \geq 0 \)) and that \( 0 \leq \theta \leq 1 \). Because hazard rates can be larger than one in continuous time, we do not impose \( p_1, q_1, \) and \( q_2 \leq 1 \). As to \( w \), we impose \( 0.01\% \leq w \leq 1 \), choosing a very small but positive lower bound so the model itself ensures the “seeding” of the contagion process among imitators.

Table 2 reports, for tetracycline, estimates obtained through direct integration (DI) for IIM, PTM and MIM as well as those from the Srinivasan-Mason (SM) procedure for PTM and MIM. Clearly, both procedures produce very similar estimates, though direct integration has somewhat higher serial correlation because it fits the cumulative adoptions \( X(t) \) rather than the periodic adoptions \( X(t) - X(t-1) \). The difference in dependent variable also explains why direct integration produces higher \( R^2 \) values. The parameter estimates of the IIM, with the zero value of \( q_1 \) meaning that segment 1 consists of independents and the high value of \( \theta \) meaning that contagion affected only a minority, are consistent with previous analyses using individual-level data on adoption times and actual network structure (Coleman et al. 1966; Van den Bulte and Lilien 2003), and so is the decomposition of total adoptions in Figure 5 (from PTM-SM re-estimated without serial correlation). The graph indicates that by month 11, when 25% of all physicians still had to adopt, all imitators had already adopted and the “laggards” consisted only of independents. This is consistent with the original finding by Coleman et al. (1966) using individual-level data that the laggards tended to be very poorly integrated in the social network and hence unaffected by social influence. Finally, the mixture models generate an estimate of \( M \)
close to the entire sample of physicians (N = 125), whereas the mixed-influence estimates are
very close to the number of adopters having adopted at the end of the observation period (X(t_17) = 109). This is consistent with our analytical result that imposing a mixed-influence model on a
mixture process can generate the kinds of estimation artifacts documented by Van den Bulte and
Lilien (1997).

Table 2. IIM, PTM and MIM results for *Medical Innovation* tetracycline data, using
estimation by direct integration (DI) and by the Srinivasan-Mason procedure (SM) *

<table>
<thead>
<tr>
<th>Model</th>
<th>M</th>
<th>p_1</th>
<th>q_1</th>
<th>q_2</th>
<th>θ</th>
<th>w</th>
<th>AR1</th>
<th>AR2</th>
<th>DW</th>
<th>MSE</th>
<th>R^2</th>
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<tbody>
<tr>
<td>IIM-DI</td>
<td>128.2^c</td>
<td>0.100^c</td>
<td>0^*</td>
<td>1.055^b</td>
<td>0.82^{c/c}</td>
<td>0.01%^*</td>
<td>0.39</td>
<td>-</td>
<td>2.14</td>
<td>2.10</td>
<td>0.999</td>
</tr>
<tr>
<td>PTM-DI</td>
<td>128.2^c</td>
<td>0.099^c</td>
<td>-</td>
<td>1.054^b</td>
<td>0.82^{c/c}</td>
<td>0.01%^*</td>
<td>0.39</td>
<td>-</td>
<td>2.14</td>
<td>2.10</td>
<td>0.999</td>
</tr>
<tr>
<td>PTM-SM</td>
<td>131.2^c</td>
<td>0.097^c</td>
<td>-</td>
<td>1.059^c</td>
<td>0.81^{c/c}</td>
<td>0.03^{c/c}</td>
<td>0.09</td>
<td>-</td>
<td>1.80</td>
<td>2.20</td>
<td>0.908</td>
</tr>
<tr>
<td>MIM-DI</td>
<td>112.6^c</td>
<td>0.092^b</td>
<td>0.170</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>0.11</td>
<td>0.11</td>
<td>1.54</td>
<td>4.72</td>
<td>0.996</td>
</tr>
<tr>
<td>MIM-SM</td>
<td>111.3^c</td>
<td>0.085^c</td>
<td>0.188^a</td>
<td>-</td>
<td>-</td>
<td>-</td>
<td>0.32</td>
<td>-</td>
<td>1.82</td>
<td>4.35</td>
<td>0.784</td>
</tr>
</tbody>
</table>

* AR1, AR2 = first-order and second-order serial correlation, DW = Durbin-Watson statistic
* Boundary constraint.
\[ p \leq .05, \quad p \leq .01, \quad p \leq .001 \]; for θ and w, the entry left of the slash (/) refers to the significance of the test against
0 and those to the right refer to the test against 1.

Figure 5. Actual and predicted number of adopters in *Medical Innovation*
Table 3 reports the results of estimating the IIM to all 33 data sets\textsuperscript{12}. Values for $p_1$ tend to be smaller than 0.3. There are two exceptions to this: Foreign Language where $\theta$ is so low that $f_m(0) = \theta p_1$ equals only 0.04, and the Beastie Boys CD that exhibited an extreme “blockbuster” pattern, i.e., extremely quickly declining sales. Values for $q_1$ show much more variance. This is especially so for CDs. For about half of them, $q_1$ equals zero, indicating the absence of word-of-mouth among influentials. In four cases, $q_1$ is larger than one, suggesting very strong word-of-mouth among influentials. However, these large estimates are very imprecise and only one is significant at 95% confidence. Values for $q_2$ also show considerable variance, with several high values recorded for the set of miscellaneous innovations. These high values may result from the strong left skew in the adoption time series without implying the presence of true contagion (Bemmaor and Lee 2002). Finally, $\theta$ is often significantly different from both 0 and 1, indicating that the IIM does not reduce to the mixed-influence or logistic models, and only weakly correlated with $w (r = -.14)$. That $\theta$ is often larger than 2.5% or 16%, traditional values used to separate innovators from imitators based on time of adoption, is an indication—in addition to the $\phi(t)$ function and the results for Medical Innovation—that the dichotomy based on drivers of adoption underlying the IIM is conceptually different from that based on time of adoption. Since it may be of particular interest, we show in Figure 6 that the IIM can indeed capture bimodal patterns in real data.

5.3. Descriptive performance

To assess the descriptive performance of the two-segment model, we compare it against that of the mixed-influence (MIM), Gamma/Shifted Gompertz (G/SG), Weibull-Gamma (WG), and Karmeshu-Goswami (KG) models. Since all these models have a closed-form solution, we

\textsuperscript{12} We do not report the ceiling parameter values $M$ due to space constraints.
estimate them using the standard Srinivasan-Mason (1986) approach. Similarly, when $q_1$ in the IIM was estimated to be zero, we estimate the closed-form solution of the PTM using the same method. However, when $q_1$ was estimated to be greater than zero, we estimate the IIM using
direct integration, difference the predicted cumulative values to obtain the predicted adoptions in each time period, and then compute the total sum of squared errors (SSE) against the actual number of adoptions. Note, this indirect approach to compute the SSE-in-adoptions can not favor the two-segment model’s performance compared to that of other models estimated directly with the goal to minimize the SSE-in-adoptions. Also, we imposed \( X(0) = 0 \) in the direct integration approach, corresponding to the \( F(0) = 0 \) assumption in the closed-form solutions used in the Srinivasan-Mason approach. In this assessment of descriptive performance, we estimate all models (including the IIM) without serial correlation. This gives a more informative assessment of the models’ descriptive performance because incorporating serial correlation into a model might alleviate a poor fit of its mean function to the data (Franses 2002).

As measures of descriptive performance, we use the mean square error (MSE) and Bayesian Information Criterion (BIC). The MSE is computed as the SSE divided by the error degrees of freedom, i.e., the number of observations minus the number of free parameters, and so penalizes models with a larger number of free parameters. We compute the BIC using the relation between
the SSE and the concentrated log-likelihood which allows one to compute likelihood ratios and other statistics based on the model likelihood from the SSE (Seber and Wild 1989). To aid interpretation, we report only the ratio of the baseline models’ MSE to that of the two-segment model. This relative measure controls for the total variance in the data, with 1 being the neutral value and higher values indicating superior fit of the two-segment model. Similarly, we report only the difference in BIC, with 0 being the neutral value and higher values indicating superior fit of the two-segment model.

Table 4 reports the performance indicators averaged for each of the four sets of data as well as for all 33 data series (Appendix A4 reports results for the individual series). The MSE ratios indicate that the two-segment model fits markedly better than the MIM, G/SG and WG models, the latter having an MSE that is 60% to 96% higher. Importantly, the model fits better for tetracycline, high-tech, and music CDs, but not for the miscellaneous products where the presumption of a discrete mixture is not strong a priori. The two-segment model outperforms the continuous-mixture KG model by only a moderate margin for tetracycline and the high-tech products. The same set of conclusions flow from using the BIC difference, where a 3-point difference is large enough to be evidence of superior fit and a 10-point difference provides strong to very strong evidence of superior fit (Raftery 1995). The one difference between the BIC and MSE analysis is that the G/SG and WG models do not underperform as much on the BIC as on the MSE for the high-tech products.
Table 4. Descriptive performance of two-segment models compared to mixed-influence, Gamma/Shifted Gompertz, Weibull-Gamma, and Karmeshu-Goswami models

<table>
<thead>
<tr>
<th></th>
<th>MSE ratio</th>
<th>BIC difference</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>MIM</td>
<td>G/SG</td>
</tr>
<tr>
<td>Tetracycline</td>
<td>2.06</td>
<td>2.22</td>
</tr>
<tr>
<td>Music CDs</td>
<td>2.27</td>
<td>1.98</td>
</tr>
<tr>
<td>High-tech</td>
<td>2.18</td>
<td>1.42</td>
</tr>
<tr>
<td>Miscellaneous</td>
<td>1.30</td>
<td>1.00</td>
</tr>
<tr>
<td>All</td>
<td>1.96</td>
<td>1.59</td>
</tr>
</tbody>
</table>

# To save space and aid interpretation, we report only the alternative models’ MSE divided by that of the two-segment model, and the alternative models’ BIC minus that of the two-segment model. For the MSE ratio, the neutral value is 1; for the BIC difference, it is 0. For both metrics, higher values indicate superior fit of the two-segment model. For the MSE ratio, the geometric mean is reported as this is a better measure of central tendency of a ratio than the arithmetic mean. For the BIC difference, the latter is reported.

6. Conclusion

We have modeled the diffusion of innovations in a social structure that is of increasing interest to marketing practitioners as well as academics: a two-segment structure with influentials who are more in touch with new developments, who may but need not act independently from each other, and who affect another segment of pure imitators whose own adoptions do not affect the influentials. Such a two-segment structure is consistent with several theories in sociology and diffusion research, including the classic two-step flow hypothesis and Moore’s more recent technology adoption framework. Our model allows diffusion researchers to operationalize these theories without recourse to micro-level diffusion data and to estimate parameters from real data. There are four main results.

1. Diffusion in a mixture of influentials and imitators can exhibit the traditional symmetric-around-the-peak bell shape, asymmetric bell shapes, as well as a dip or “chasm” between the early and later parts of the diffusion curve. In contrast to Moore’s contention, the model suggests that it need not always be necessary to change the product to gain traction among later adopters
and the adoption curve to swing up again. Tetracycline is an example. Hence, we provide closed-
form corroboration for the result Goldenberg et al. (2002) obtained from a simulation study. The
management implication is that launching a new version to appeal to prospects who have not
adopted yet need not be necessary, let alone optimal, to get out of the dip.

(2) The proportion of adoptions stemming from independents need not decrease
monotonically; it can also first decline and then rise again to unity. This rejects a still common
contention based on an erroneous mixture interpretation of the mixed-influence model. The
management implication is that, while it may make sense to shift the focus of one’s marketing
efforts from independents to imitators shortly after launch as shown by Mahajan and Muller
(1998) using a two-period model, one may want to start increasing one’s resource allocation to
independent decision makers again later in the process.

(3) Specifying a mixed-influence model to a mixture process where influentials act
independently from each other \(q_1=0\) can generate systematic changes in the parameter values.
As several authors have noted, diffusion within a pure-type mixture of independents and
imitators with hazards \(p\) and \(qF(t)\), respectively, is distinct from diffusion in a homogenous
population with mixed-influence where everyone adopts with hazard \(p + qF(t)\). The closed-form
solutions we present not only prove this mathematically but also show that unless \(0\) is close to
either 0 or 1, imposing a mixed-influence specification on a pure-type mixture process can
generate the systematic changes in the parameter values reported by Van den Bulte and Lilien

(4) Empirical analysis of four sets of data comprising a total of 33 different data series (the
classic Medical Innovation data, 19 music CDs, 5 high-tech products, and 8 miscellaneous
innovations) indicates that the two-segment model fits markedly better than the mixed-influence,
the Gamma/Shifted Gompertz, and the Weibull-Gamma model, at least in cases where a two-
segment structure is likely to exist. Hence, the model does better when it is theoretically
expected to and does not when it is not theoretically expected to. The two-segment model fits
only somewhat better, but certainly not worse, than the mixed-influence model recently proposed
by Karmeshu and Goswami (2001) where \( p \) and \( q \) vary in a continuous fashion. Specifically, the
two-segment model outperforms the continuous-mixture KG model by a moderate margin in two
of the four cases (tetracycline and the high-tech products) and matches it in the other two cases
(music CDs and miscellaneous innovations). Overall, the differences in descriptive performance
indicate that the discrete-mixture model is sufficiently different and the data sufficiently
informative for the model to fit real data better than other models.

The models we presented provide sharper insight into how social structure can affect macro-
level diffusion patterns, and should prove useful in three areas of application where influentials
and imitators are a priori likely to exist. The first area is that of high-technology products, where
there is a strong interest in the potential existence of “early market” independents and
“mainstream” imitators. The second area is that of movies, a category that has attracted the
attention of many researchers recently.\(^{13}\) The third area consists of situations where a segment of
enthusiasts has pent-up demand. For instance, when internet access providers started operating in
France in 1996, a rather large number of people adopted their services. New adoptions dipped in

\[^{13}\text{Explicitly allowing for influentials and imitators may be especially useful for movies carried by actors or directors who already have a small following among aficionados but have not yet broken through to the mainstream audience. In such cases, one would expect the former to adopt according an independent process and the latter to adopt only through contagion, if at all. This might result in a temporary dip. Movies starring Christina Ricci and movies directed by Ang Lee exhibit this pattern. Early in her career, Christina Ricci played in several independent movies that won critical acclaim and earned her the label of “Indie Queen”. These early movies exhibited the bell curve typical of very successful “sleepers” (The Ice Storm-1997; The Opposite of Sex-1998; Buffalo 66-1998). Then followed a small movie exhibiting a dip (Desert Blue-1998), while her recent movies are more standard Hollywood fare exhibiting the standard monotonic, exponential decline (e.g., The Man Who Cried-2001). The same pattern is observed for movies directed by Ang Lee; bell-shaped for The Ice Storm-1997, a temporary dip for Ride with the Devil-1999, and monotonic decline for his more recent Hollywood production The Hulk-2003.\]
1997, only to increase again from 1998 onwards. The deviation from the standard bell shape was not the low number in 1997 but the high initial number in 1996, when many university users who had been accessing the internet exclusively through the university RENATER network were finally able to start using the internet at home as well (Fornerino 2003). In case the enthusiasts can place advance orders that the marketing analyst can observe (e.g., Moe and Fader 2002), it may be useful to explicitly allow for a difference between the start time of the diffusion process of the two segments.

An important extension of our work would be to incorporate control variables, including marketing mix efforts. This may not only be useful for empirical research, but may also enable one to study the decision to target independents versus imitators in continuous time. Even a simplified three-period model might be helpful in studying under what conditions it is profit maximizing to change one’s targeting from independents to imitators and, possibly, to independents again (cf. Mahajan and Muller 1998). Like the models we presented, this extension would allow one to better understand current arguments and findings, to formalize richer theoretical arguments, and perhaps even to operationalize them into estimable models that help bridge the gap between theory and data.
References


Appendix

A.1. Solution for $F_2(t)$ in IIM: $h_1(t) = p_1 + q_1 F_1(t)$ and $h_2(t) = q_2[wF_1(t)+(1-w)F_2(t)]$

To simplify notation, we omit the time argument from functions and write $h_1$ instead of $h_1(t)$, etc. We know $F_1 = (1 - e^{-(p_1+q_1)}(t))/(1 + q_1 e^{-(p_1+q_1)}(t))$. Since $h_2 = q_2(wF_1 + (1 - w)F_2)$, we write:

$$\frac{dF_2}{dt} = q_2 w \frac{1 - e^{-(p_1+q_1)}(t)}{1 + q_1 e^{-(p_1+q_1)}(t)} + q_2 (1 - w - w \frac{1 - e^{-(p_1+q_1)}(t)}{1 + q_1 e^{-(p_1+q_1)}(t)}) F_2 - q_2 (1 - w) F_2^2$$

[A.1.1]

A.1.1 is a Ricatti equation of the general form $\frac{dy}{dx} = P(x) + Q(x)y + R(x)y^2$.

Setting $y = F_2$ and $x = t$, we get $P(x) = q_2 w \frac{1 - e^{-(p_1+q_1)}(t)}{1 + q_1 e^{-(p_1+q_1)}(t)}$; $Q(x) = q_2 (1 - w - \frac{1 - e^{-(p_1+q_1)}(t)}{1 + q_1 e^{-(p_1+q_1)}(t)})$; and

$R(x) = -q_2 (1 - w)$. $F_2 = 1$ is a potential solution for this Ricatti equation. We use the transformation $z = \frac{1}{F_2 - 1} \Rightarrow F_2 = \frac{z + 1}{z}$. For $F_2$ continuous in $[0,1]$, $z$ is continuous in $(-\infty, -1]$.

Note, $\frac{dF_2}{dt} = -\frac{1}{z^2} \frac{dz}{dt}$. The equation now becomes:

$$\frac{dz}{dt} = q_2 (1 - w) + q_2 z (1 - w + w \frac{1 - e^{-(p_1+q_1)}(t)}{1 + q_1 e^{-(p_1+q_1)}})$$

[A.1.2]

This is of the form $\frac{du}{dx} + P_1(x) \mu = Q_1(x)$ with $\mu = z$; $x = t$; $P_1(x) = -q_2 (1 - w + w \frac{1 - e^{-(p_1+q_1)}(t)}{1 + q_1 e^{-(p_1+q_1)}})$; and

$Q_1(x) = q_2 (1 - w)$. The general solution for such an equation is $\mu = \frac{\int u(x)Q_1(x)dx + c}{u(x)}$ where $u(x) = \exp(\int P_1(x)dx)$ is the integrating factor.

Since $\int P_1(x)dx = -q_2 t - \frac{q_2 \ln(p_1 + q_1 e^{-(p_1+q_1)t})}{q_1}$, we get $u(x) = \exp(-q_2 t - \frac{q_2 \ln(p_1 + q_1 e^{-(p_1+q_1)t})}{q_1})$. 

A-1
Hence, \( \int u(x)Q_1(x)\,dx = \frac{e^{-\beta(x_1,\lambda_2,\lambda_1,\lambda_2,\lambda_1)}}{p_{\lambda_2-q_{\lambda_1}}(1-w)} \), where

\[ H_1 = 2F_1 \left(1, \frac{w}{q_1}; 1 + \frac{w}{q_1}; \frac{q_{1}}{p_{1} + q_{1}} ; \frac{p_{1}}{p_{1} + q_{1} + q_{1} e^{(-p_{1} + q_{1})}} \right) \] and \( 2F_1(1, b; c; k) \) is the Gaussian hypergeometric function, the series representation of which is given by

\[ \sum_{n=0}^{\infty} \frac{\Gamma(b+n)\Gamma(c)}{\Gamma(b)\Gamma(c+n)} k^n . \] This series is convergent for arbitrary \( b, c \) when \( |k| < 1 \); and for \( k = \pm 1 \) when \( c > 1 + b \). This implies that the series is convergent as long as \( q_1 > 0 \).

Substituting back, we get

\[ z = \frac{e^{-\beta(x_1,\lambda_2,\lambda_1,\lambda_2,\lambda_1)}}{p_{\lambda_2-q_{\lambda_1}}(1-w)} + c \]

Transforming \( z \) back to \( F_2 \), we obtain

\[ F_2 = 1 + \frac{e^{-\beta(x_1,\lambda_2,\lambda_1,\lambda_2,\lambda_1)}}{p_{\lambda_2-q_{\lambda_1}}(1-w)} \frac{\ln(p_{1} + q_{1} e^{(-p_{1} + q_{1})})}{q_{1}} + c \] \[ \text{[A.1.3]} \]

Since \( F_2(0) = 0 \), \( c = \frac{(p_{1} + q_{1})}{p_{\lambda_2-q_{\lambda_1}}(1-w)} \), where

\[ H_2 = 2F_1 \left(1, \frac{w}{q_1}; 1 + \frac{w}{q_1}; \frac{q_{1}}{p_{1} + q_{1}} ; \frac{p_{1}}{p_{1} + q_{1}} \right) \]. Simplifying, we obtain as closed-form expression:

\[ F_2(t) = 1 + \frac{p_{1} w - q_{1} (1-w)}{q_{1} (1-w)H_1 + e^{\beta(x_1,\lambda_2,\lambda_1,\lambda_2,\lambda_1)} \left( \frac{p_{1} + q_{1} e^{(-p_{1} + q_{1})}}{p_{1} + q_{1}} \right) (q_{1} (1-w) - p_{1} w - q_{1} (1-w)H_2) \] \[ \text{[A.1.4]} \]

A.2. Solution for \( F_2(t) \) in PTM: \( h_1(t) = p_{1} \) and \( h_2(t) = q_{2} \left[ wF_1(t) + (1-w)F_2(t) \right] \)

Note: to simplify the notation, we replace \( p_{1} \) by \( p \) and \( q_{2} \) by \( q \). We know that \( F_1 = 1 - e^{-\beta t} \) and that \( f_2 = \frac{dF_2}{dt} = q[\left[ wF_1(t) + (1-w)F_2(t) \right][1 - F_2(t)] \), and so write

\[ \frac{dF_2}{dt} = qw(1-e^{-\beta t}) + q(1-2w+we^{-\beta t})F_2 - q(1-w)F_2^2 \] \[ \text{[A.2.1]} \]
The above equation is a Ricatti equation of the general form:

\[ \frac{dy}{dx} = P(x) + Q(x)y + R(x)y^2 \]  \[ \text{[A.2.2]} \]

Setting \( y = F_2 \) and \( x = t \), we get \( P(x) = qw(1 - e^{-pt}) \); \( Q(x) = q(1 - 2w + we^{-pt}) \); and \( R(x) = -q(1 - w) \). \( F_2 = 1 \) is a potential solution for Eq[A.2.2]. We use the transformation

\[ z = \frac{1}{F_2 - 1} \Rightarrow F_2 = z + 1 \]

For \( F_2 \) continuous in \([0,1]\), \( z \) is continuous in \((-\infty,-1]\). Note

\[ \frac{dF_2}{dt} = -\frac{1}{z^2} \frac{dz}{dt} \]

After substitution, we can write Eq[A.2.1] as

\[ \frac{dz}{dt} = q(1 - w) + zq(1 - we^{-pt}) \]  \[ \text{[A.2.3]} \]

Eq[A.2.3] is of the form:

\[ \frac{d\mu}{dx} + P_1(x)\mu = Q_1(x) \]

with \( \mu = z; \ x = t; \ P_1(x) = -q(1 - we^{-pt}) \); and \( Q_1(x) = q(1 - w) \). Again, the general solution for such an equation is

\[ \mu = \int u(x)Q_1(x)dx + c \]

where \( u(x) = \exp(\int P_1(x)dx) \) is the integrating factor.

Since \( \int P_1(x)dx = -qt + \frac{qwe^{-pt}}{-p} \), we get \( u(x) = \exp(-qt - \frac{q}{p} we^{-pt}) \); and hence

\[ \int u(x)Q_1(x)dx = \int \exp(-qt - \frac{q}{p} we^{-pt})q(1 - w)dt \]

Now let us define

\[ I = q(1 - w)\int \exp(-qt - \frac{q}{p} we^{-pt})dt \]  \[ \text{[A.2.5]} \]

To solve this integral, we do another transformation: \( a = e^{-pt} \Rightarrow t = -\frac{1}{p}\ln a \Rightarrow dt = -\frac{1}{pa} da \).

Eq[A.2.5] then becomes

\[ I = -\frac{q}{p} (1 - w)\int a^{\frac{q}{p}-1} \exp(-\frac{q}{p} wa)da \]

The solution to which is

\[ I = \frac{q}{p} (1 - w)(\frac{q}{p} w)^{\frac{q}{p}-1} \Gamma(\frac{q}{p}, \frac{q}{p} wa) \]  \[ \text{[A.2.6]} \]

where \( \Gamma(\eta,k) \) is the “upper” incomplete gamma function: \( \Gamma(\eta,k) = \int_k^{\infty} e^{-v} dv \).
Substituting \( a = e^{-pt} \) in Eq[A.2.6], and then \( I = \int u(x)Q_t(x)dx \) from Eq[A.2.6] back in Eq[A.2.4],

\[
z(t) = \frac{I + c}{u(x)} = \frac{\frac{q}{p}(1-w)(\frac{q}{p}w)^{-\frac{1}{2}}\Gamma(\frac{q}{p}, \frac{q}{p}we^{-pt}) + c}{\exp(-qt - \frac{q}{p}we^{-pt})}
\]  

[A.2.7]

Transforming \( z \) back to \( F_2 \), we get:

\[
F_2(t) = 1 + \frac{\exp(-qt - \frac{q}{p}we^{-pt})}{\frac{q}{p}(1-w)(\frac{q}{p}w)^{\frac{1}{2}}\Gamma(\frac{q}{p}, \frac{q}{p}we^{-pt}) + c}
\]

[A.2.8]

Since \( F_2(0) = 0 \), we get \( c = -\exp(-\frac{q}{p}w) - \frac{q}{p}(1-w)(\frac{q}{p}w)^{\frac{1}{2}}\Gamma(\frac{q}{p}, \frac{q}{p}w) \). Hence:

\[
F_2(t) = 1 + \frac{\exp(-qt - \frac{q}{p}we^{-pt})}{\frac{q}{p}(1-w)(\frac{q}{p}w)^{\frac{1}{2}}(\Gamma(\frac{q}{p}, \frac{q}{p}we^{-pt}) - \Gamma(\frac{q}{p}, \frac{q}{p}w)) - \exp(-\frac{q}{p}w)}
\]

[A.2.9]

### A.3. Relation of PTM3 to Jeuland’s (1981) analysis

Jeuland (1981) defines \( Y_S(T) \) and \( Y_I(T) \) as the cumulative number of adoptions by pure independents and pure imitators, respectively. Similarly, \( m_S \) and \( m_I \) are the number of adopters of each type who will ultimately adopt. He posits that adoptions by imitators follow the following equation (Jeuland 1981, p. 12, eq. 19):

\[
\frac{dY_I(T)}{dT} = \beta_1[m_I - Y_S(T)][Y_S(T) + Y_I(T)],
\]

which he then rewrites as (Jeuland 1981, p. 12, eq. 20):

\[
\frac{dY_I(T)}{dT} = \beta_1m_SY_S(T) + \beta_1(m_I - Y_S(T))Y_I(T) - \beta_1Y_I^2(T)
\]

[A.3.1]

Jeuland's (1981, p. 14) proposed solution for the above equation is

\[
Y_I(T) = m_I(1 - \frac{g(T)}{1 - \beta_1m_Ig(T)})
\]

[A.3.2]

where \( g(T) = \exp[-\beta_1(m_I T + YC_S(T))] \); \( YC_S(T) = \int_0^T Y_S(v)dv \); \( G(T) = \int_0^T g(u)du \).

When \( Y_S(T) \) follows an exponential process with \( dY_S(T)/dT = p[m_S - Y_S(T)] \), our notation has the following identities with Jeuland's notation:
\[ Y_S(T) = \theta \eta (1 - e^{-pT}) \]
\[ Y_f(T) = (1 - \theta)mF_z(T) \]
\[ \beta_f = \frac{q_2}{m} \]
\[ m_y = (1 - \theta)m \]
\[ m_s = \theta \eta \]

Again, to simplify the notation, we replace \( p_1 \) by \( p \) and \( q_2 \) by \( q \).

Substituting into Eq[A.3.1] and simplifying gives us a Ricatti equation identical to Eq[A.2.1], apart from substituting \( \theta \) for \( w \).

We next derive the expressions for \( YC_S(T), g(T), \) and \( G(T) \).

\[
YC_S(T) = \int_0^T Y_S(v)dv = \int_0^T \theta \eta (1 - e^{-pv})dv = \frac{\theta \eta}{p} (-1 + pT + e^{-pT}) \tag{A.3.3}
\]

\[
g(T) = \exp[-\beta_f(m_yT + YC_S(T))] = \exp[-qT + \frac{q\theta}{p} (1 - e^{-pT})] \tag{A.3.4}
\]

\[
G(T) = \int_0^T g(u)du = \int_0^T \exp[-qu + \frac{q\theta}{p} (1 - e^{-pu})]du = e^{\frac{q}{p}} \int_0^T \exp[-qu - \frac{q\theta}{p} e^{-pu}]du \tag{A.3.5}
\]

Eq[A.3.5] is similar in form to Eq[A.2.5]. We solve it using the same transformation and obtain:

\[
G(T) = \frac{1}{p} \exp\left[\frac{q}{p} \theta\right] \left(\frac{q}{p} \theta\right)^{-\frac{q}{p}} \left(\Gamma\left(\frac{q}{p}, \frac{q}{p} e^{-pT}\right) - \Gamma\left(\frac{q}{p}, \frac{q}{p} \theta\right)\right) \tag{A.3.6}
\]

Substituting Eq[A.3.4] and Eq[A.3.6] into Eq[A.3.2] and simplifying, one obtains:

\[
Y_f(T) = m_y (1 + \frac{\exp(-qT - \frac{q}{p} \theta e^{-pT})}{\frac{q}{p} (1 - \theta)(\frac{q}{p} \theta)^{-\frac{q}{p}} \left(\Gamma\left(\frac{q}{p}, \frac{q}{p} \theta e^{-pT}\right) - \Gamma\left(\frac{q}{p}, \frac{q}{p} \theta\right)\right)} - \exp(-\frac{q}{p} \theta) \right) \tag{A.3.7}
\]

Thus, Jeuland’s proposed solution, when fully worked out for the case where \( Y_S(T) \) follows an exponential process, results in our closed-form solution to PTM3.
A.4. Relative descriptive performance for all 33 data series

<table>
<thead>
<tr>
<th></th>
<th>MSE ratio</th>
<th>BIC difference</th>
</tr>
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<tbody>
<tr>
<td></td>
<td>MIM G/SG</td>
<td>WG KG</td>
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<tr>
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<td>2.22 2.38 1.36</td>
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<td>2.42 4.33 0.96</td>
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<td>4.23 4.23 0.98</td>
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<td>4.12</td>
<td>3.17 4.73 1.73</td>
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<td>3.33</td>
<td>3.33 0.95 0.99</td>
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<tr>
<td>Bonnie Raitt 2</td>
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<td>2.83 6.34 1.64</td>
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<td>2.02 1.68 0.61</td>
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<td>2.89 2.24 0.82</td>
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<td>1.43 1.61 0.77</td>
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<td>0.99 - 1.16</td>
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<tr>
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<td>1.22 1.12 1.00</td>
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<td>1.23 1.28 0.71</td>
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<td>Luscious Jackson</td>
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<td>1.45 - 0.95</td>
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<td>1.55 1.00 0.94</td>
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<tr>
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<td>1.06 1.05 1.08</td>
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</tbody>
</table>

Note: To save space and aid interpretation, we report only the alternative models’ MSE divided by that of the two-segment model, and the alternative models’ BIC minus that of the two-segment model. For the MSE ratio, the neutral value is 1; for the BIC difference, it is 0. For both metrics, higher values indicate superior fit of the two-segment model. The entry “-” indicates that the model did not converge.